

Online Appendix

1 Additional Tables

Table A1: Summary Statistics

| | Median | Mean | St. Dev. | St. Dev. w/in Countries | Min. | Max. |
|--|--------|---------|----------|-------------------------|---------|---------|
| Democratization | 0 | .019 | .137 | .139 | 0 | 1 |
| Authoritarian Breakd. | 0 | .078 | .268 | .26 | 0 | 1 |
| Democratization after Breakd. | 0 | .249 | .433 | .343 | 0 | 1 |
| % Dem. neigh. (spatial lag 1) | .167 | .218 | .257 | .189 | 0 | 1 |
| Ch. % dem. neigh. last year (spatial lag 1) | 0 | .004 | .076 | .076 | -1 | 1 |
| Ch. % dem. neigh. last 2 years (spatial lag 1) | 0 | .008 | .106 | .105 | -1 | 1 |
| Ch. % dem. neigh. last 3 years (spatial lag 1) | 0 | .012 | .129 | .126 | -1 | 1 |
| % Dem. world | .321 | .341 | .119 | .093 | .132 | .584 |
| Ch. % dem. world last year | 0 | .004 | .016 | .016 | -.035 | .079 |
| Ch. % dem. world last 2 years | .002 | .008 | .026 | .025 | -.057 | .101 |
| Ch. % dem. world last 3 years | .005 | .011 | .035 | .037 | -.062 | .13 |
| % Dem. neigh. (spatial lag 2) | .2 | .248 | .21 | .155 | 0 | 1 |
| Ch. % dem. neigh. last year (spatial lag 2) | 0 | .004 | .047 | .047 | -.455 | .667 |
| Ch. % dem. neigh. last 2 years (spatial lag 2) | 0 | .008 | .066 | .065 | -.5 | .667 |
| Ch. % dem. neigh. last 3 years (spatial lag 2) | 0 | .012 | .081 | .079 | -.5 | .667 |
| % Dem. region | .143 | .192 | .221 | .147 | 0 | .9 |
| Ch. % dem. region last year | 0 | .003 | .042 | .042 | -.333 | .5 |
| Ch. % dem. region last 2 years | 0 | .006 | .058 | .058 | -.5 | .5 |
| Ch. % dem. region last 3 years | 0 | .008 | .07 | .069 | -.5 | .5 |
| Ch. # dem. neigh. last year (spatial lag 1) | 0 | .023 | .358 | .353 | -6 | 3 |
| Ch. # dem. neigh. last year (spatial lag 2) | 0 | .081 | .772 | .762 | -13 | 16 |
| Hegemonic power volatility | .411 | .604 | .518 | .486 | .057 | 3.203 |
| Growth rate | 1.025 | 1.63 | 5.417 | 5.161 | -33.291 | 158.606 |
| Growth crisis | 0 | .331 | .471 | .393 | 0 | 1 |
| Bench. Growth | -.284 | -.136 | 5.304 | 5.042 | -34.306 | 155.251 |
| Rain dev. | -.08 | -.076 | 1.903 | 1.833 | -12.623 | 14.508 |
| RR crisis | 1 | .523 | .5 | .462 | 0 | 1 |
| LV crisis | 1 | .564 | .496 | .416 | 0 | 1 |
| RR/LV crisis | 1 | .557 | .497 | .445 | 0 | 1 |
| GDP pc (logged) | 7.507 | 7.609 | .788 | .407 | 5.093 | 10.758 |
| Oil | 0 | 440.786 | 3132.657 | 2050.923 | 0 | 78588.8 |
| Muslim | 1 | 24.136 | 37.353 | 0 | 0 | 99.7 |
| Catholic | 31 | 42.55 | 40.651 | 0 | 0 | 99 |
| Protestant | 2 | 9.42 | 17.104 | 0 | 0 | 90.6 |
| Ethnic frac. | 27 | 30.875 | 24.561 | 0 | 0 | 82 |
| Rel. frac. | 13 | 22.849 | 21.761 | 0 | 0 | 83.5 |
| British Col. | 0 | .154 | .361 | 0 | 0 | 1 |
| # Past Break. | 0 | .221 | .542 | .353 | 0 | 4 |
| Polarized Order | 1 | .578 | .494 | .429 | 0 | 1 |
| Democratic Order | 0 | .192 | .394 | .367 | 0 | 1 |
| Alliance with France | 0 | .012 | .109 | .103 | 0 | 1 |
| Alliance with Germany | 0 | .012 | .111 | .095 | 0 | 1 |
| Alliance with Japan | 0 | .003 | .052 | .05 | 0 | 1 |
| Alliance with Russia/USSR | 0 | .054 | .225 | .18 | 0 | 1 |
| Alliance with US | 0 | .134 | .341 | .274 | 0 | 1 |
| Alliance with US (Cold War) | 0 | .113 | .317 | .262 | 0 | 1 |
| Alliance with UK | 0 | .019 | .135 | .128 | 0 | 1 |
| Pro-Democratic Hegemonic Shocks | 0 | .137 | .344 | .337 | 0 | 1 |

Table A2: Cases of Authoritarian Breakdown — Part I

| | | | | |
|----------------------------|-------------------------|-------------------------|-------------------------|-------------------------|
| Ecuador 1875 | Iran 1896 | Dominican Rep 1914 | El Salvador 1931 | Paraguay 1948 |
| <i>Uruguay 1875</i> | Netherlands 1897 | Haiti 1914 | Panama 1931 | Peru 1948 |
| Bolivia 1876 | Spain 1897 | Mexico 1914 | Peru 1931 | Venezuela 1948 |
| <i>Dominican Rep 1876</i> | Uruguay 1897 | Peru 1914 | Chile 1932 | China 1949 |
| Ecuador 1876 | El Salvador 1898 | Haiti 1915 | Ecuador 1932 | <i>Indonesia 1949</i> |
| El Salvador 1876 | Guatemala 1898 | Dominican Rep 1916 | <i>Japan 1932</i> | Panama 1949 |
| Haiti 1876 | Venezuela 1898 | <i>Ethiopia 1916</i> | <i>Afghanistan 1933</i> | Paraguay 1949 |
| <i>Honduras 1876</i> | Bolivia 1899 | China 1917 | <i>Cuba 1933</i> | Syria 1949 |
| Mexico 1876 | Dominican Rep 1899 | <i>Costa Rica 1917</i> | Ecuador 1933 | Haiti 1950 |
| Turkey 1876 | Venezuela 1899 | Greece 1917 | Peru 1933 | Pakistan 1950 |
| <i>Uruguay 1876</i> | Colombia 1900 | <i>Russia 1917</i> | Austria 1934 | Panama 1950 |
| Paraguay 1877 | Denmark 1901 | Germany 1918 | Bolivia 1934 | Taiwan 1950 |
| <i>Dominican Rep 1878</i> | <i>Nepal 1901</i> | Turkey 1918 | Chile 1934 | Venezuela 1950 |
| <i>Afghanistan 1879</i> | Dominican Rep 1902 | <i>Afghanistan 1919</i> | <i>Cuba 1934</i> | Bolivia 1951 |
| Bolivia 1879 | Haiti 1902 | <i>Costa Rica 1919</i> | Yugoslavia 1934 | Iran 1951 |
| <i>Dominican Rep 1879</i> | Paraguay 1902 | Honduras 1919 | Ecuador 1935 | Jordan 1951 |
| Haiti 1879 | Dominican Rep 1903 | Italy 1919 | Bolivia 1936 | Nepal 1951 |
| Peru 1879 | Honduras 1903 | <i>Peru 1919</i> | Paraguay 1936 | Bolivia 1952 |
| Venezuela 1879 | Paraguay 1904 | Spain 1919 | Bolivia 1937 | Egypt 1952 |
| <i>Afghanistan 1880</i> | Dominican Rep 1905 | Uruguay 1919 | China 1937 | Japan 1952 |
| Paraguay 1880 | Paraguay 1905 | Bolivia 1920 | Colombia 1937 | Panama 1952 |
| Peru 1881 | Cuba 1906 | Guatemala 1920 | Ecuador 1937 | <i>Afghanistan 1953</i> |
| Russia 1881 | Ecuador 1906 | Mexico 1920 | Paraguay 1937 | Colombia 1953 |
| Peru 1882 | Paraguay 1906 | Guatemala 1921 | Bolivia 1939 | Iran 1953 |
| <i>Uruguay 1882</i> | Honduras 1907 | Japan 1921 | Spain 1939 | Egypt 1954 |
| Ecuador 1883 | <i>Korea 1907</i> | Paraguay 1921 | Cuba 1940 | Paraguay 1954 |
| <i>Honduras 1883</i> | Haiti 1908 | Spain 1921 | <i>Estonia 1940</i> | Syria 1954 |
| <i>Dominican Rep 1885</i> | Morocco 1908 | Greece 1922 | <i>Latvia 1940</i> | Argentina 1955 |
| El Salvador 1885 | Paraguay 1908 | Ireland 1922 | <i>Lithuania 1940</i> | Indonesia 1955 |
| Guatemala 1885 | Portugal 1908 | Spain 1922 | Panama 1941 | Haiti 1956 |
| <i>Nepal 1885</i> | Venezuela 1908 | China 1923 | Uruguay 1942 | Honduras 1956 |
| Peru 1885 | Chile 1909 | Spain 1923 | Argentina 1943 | Hungary 1956 |
| United Kingdom 1885 | Nicaragua 1909 | China 1924 | Bolivia 1943 | Nicaragua 1956 |
| Uruguay 1886 | Turkey 1909 | Honduras 1924 | Italy 1943 | Peru 1956 |
| Haiti 1888 | <i>Korea 1910</i> | Ecuador 1925 | Ecuador 1944 | Colombia 1957 |
| Brazil 1889 | Nicaragua 1910 | Greece 1925 | El Salvador 1944 | Guatemala 1957 |
| <i>Costa Rica 1889</i> | Dominican Rep 1911 | Honduras 1925 | France 1944 | Haiti 1957 |
| <i>Ethiopia 1889</i> | Ecuador 1911 | Mongolia 1925 | Guatemala 1944 | Honduras 1957 |
| Haiti 1889 | Haiti 1911 | China 1926 | Brazil 1945 | Thailand 1957 |
| El Salvador 1890 | Honduras 1911 | Greece 1926 | Germany 1945 | Argentina 1958 |
| Brazil 1891 | Mexico 1911 | Nicaragua 1926 | Guatemala 1945 | Guatemala 1958 |
| Chile 1891 | Nicaragua 1911 | Chile 1927 | Venezuela 1945 | <i>Iraq 1958</i> |
| Venezuela 1892 | Paraguay 1911 | China 1928 | Bolivia 1946 | Thailand 1958 |
| <i>Honduras 1893</i> | Sweden 1911 | <i>Afghanistan 1929</i> | Haiti 1946 | Venezuela 1958 |
| <i>Nicaragua 1893</i> | Argentina 1912 | Bolivia 1930 | Philippines 1946 | <i>Cuba 1959</i> |
| Belgium 1894 | Dominican Rep 1912 | Brazil 1930 | Ecuador 1947 | El Salvador 1960 |
| El Salvador 1894 | Haiti 1912 | Guatemala 1930 | Nicaragua 1947 | Korea South 1960 |
| <i>Honduras 1894</i> | Paraguay 1912 | <i>Japan 1930</i> | Costa Rica 1948 | Laos 1960 |
| Paraguay 1894 | Spain 1912 | Peru 1930 | Ecuador 1948 | <i>Myanmar 1960</i> |
| Peru 1894 | El Salvador 1913 | Romania 1930 | El Salvador 1948 | Turkey 1960 |
| Ecuador 1895 | Mexico 1913 | Spain 1930 | <i>Indonesia 1948</i> | Dominican Rep 1961 |
| Peru 1895 | Turkey 1913 | Ecuador 1931 | Panama 1948 | El Salvador 1961 |

Note: Authoritarian breakdowns resulting in a transition to democracy are shown in **bold**. Authoritarian breakdowns shown in *italic* are omitted from our main regression (Model 1 of Table 5) due to missing data on explanatory variables.

Table A2: Cases of Authoritarian Breakdown — Part II

| | | | | |
|---------------------------|---------------------------|-------------------------|----------------------------|------------------------------|
| Syria 1961 | <i>Oman 1970</i> | <i>Yemen South 1978</i> | Uruguay 1985 | Central Afr. Rep 1993 |
| Turkey 1961 | Poland 1970 | <i>Afghanistan 1979</i> | <i>Afghanistan 1986</i> | <i>Latvia 1993</i> |
| Dominican Rep 1962 | Syria 1970 | Bolivia 1979 | Bangladesh 1986 | Madagascar 1993 |
| Syria 1962 | Argentina 1971 | Brazil 1979 | Guatemala 1986 | Niger 1993 |
| <i>Yemen North 1962</i> | Bolivia 1971 | Cambodia 1979 | Haiti 1986 | Nigeria 1993 |
| Argentina 1963 | Honduras 1971 | Central Afr. Rep 1979 | Lesotho 1986 | Burundi 1994 |
| Benin 1963 | <i>Lebanon 1971</i> | Chad 1979 | Philippines 1986 | Guinea-Bissau 1994 |
| Dominican Rep 1963 | Sudan 1971 | Congo Brazzaville 1979 | Uganda 1986 | Haiti 1994 |
| <i>Iraq 1963</i> | Uganda 1971 | Ecuador 1979 | <i>Yemen South 1986</i> | Lesotho 1994 |
| Nicaragua 1963 | Benin 1972 | El Salvador 1979 | Burkina Faso 1987 | Malawi 1994 |
| Peru 1963 | Ecuador 1972 | Equatorial Guinea 1979 | Burundi 1987 | Mozambique 1994 |
| Syria 1963 | Gambia 1972 | Ghana 1979 | Tunisia 1987 | Rwanda 1994 |
| <i>Togo 1963</i> | <i>Madagascar 1972</i> | Iran 1979 | Haiti 1988 | South Africa 1994 |
| <i>Vietnam South 1963</i> | Qatar 1972 | Korea South 1979 | Korea South 1988 | Comoros 1995 |
| Benin 1964 | <i>Afghanistan 1973</i> | Mauritania 1979 | <i>Myanmar 1988</i> | Qatar 1995 |
| Bolivia 1964 | Argentina 1973 | Nicaragua 1979 | Pakistan 1988 | <i>Afghanistan 1996</i> |
| Gabon 1964 | Greece 1973 | Nigeria 1979 | Comoros 1989 | Burundi 1996 |
| Laos 1964 | Rwanda 1973 | Uganda 1979 | Paraguay 1989 | Sierra Leone 1996 |
| Sudan 1964 | Thailand 1973 | Burkina Faso 1980 | Poland 1989 | Taiwan 1996 |
| <i>Vietnam South 1964</i> | Burkina Faso 1974 | El Salvador 1980 | Romania 1989 | Albania 1997 |
| Algeria 1965 | Cyprus 1974 | Guinea-Bissau 1980 | Bulgaria 1990 | Cambodia 1997 |
| Benin 1965 | Greece 1974 | Liberia 1980 | Chad 1990 | Congo Brazzaville 1997 |
| Congo Kinshasa 1965 | Niger 1974 | Mauritania 1980 | Chile 1990 | Congo Kinshasa 1997 |
| Dominican Rep 1965 | Portugal 1974 | Peru 1980 | Czechoslovakia 1990 | Ghana 1997 |
| <i>Vietnam South 1965</i> | <i>Yemen North 1974</i> | Uganda 1980 | <i>Germany East 1990</i> | Sierra Leone 1997 |
| Burkina Faso 1966 | Bangladesh 1975 | Argentina 1981 | Haiti 1990 | Sierra Leone 1998 |
| Burundi 1966 | Chad 1975 | Bangladesh 1981 | Hungary 1990 | Comoros 1999 |
| Central Afr. Rep 1966 | Honduras 1975 | Bolivia 1981 | Kuwait 1990 | Guinea-Bissau 1999 |
| Ecuador 1966 | Laos 1975 | Central Afr. Rep 1981 | Liberia 1990 | Indonesia 1999 |
| Ghana 1966 | <i>Madagascar 1975</i> | Egypt 1981 | Mongolia 1990 | Ivory Coast 1999 |
| Guatemala 1966 | Nigeria 1975 | Panama 1981 | Panama 1990 | Niger 1999 |
| Indonesia 1966 | Peru 1975 | Uruguay 1981 | Benin 1991 | Croatia 2000 |
| South Africa 1966 | Saudi Arabia 1975 | Argentina 1982 | <i>Estonia 1991</i> | Fiji 2000 |
| Syria 1966 | Thailand 1975 | Bangladesh 1982 | Haiti 1991 | Ivory Coast 2000 |
| Benin 1967 | <i>Vietnam South 1975</i> | Burkina Faso 1982 | Lesotho 1991 | Mexico 2000 |
| Togo 1967 | Burundi 1976 | Chad 1982 | Mali 1991 | Senegal 2000 |
| <i>Yemen North 1967</i> | Ecuador 1976 | Honduras 1982 | Nepal 1991 | <i>Yugoslavia 2000</i> |
| Benin 1968 | Nigeria 1976 | Panama 1982 | <i>Somalia 1991</i> | <i>Afghanistan 2001</i> |
| Congo Brazzaville 1968 | Portugal 1976 | Burkina Faso 1983 | Sri Lanka 1991 | Congo Kinshasa 2001 |
| Czechoslovakia 1968 | Uruguay 1976 | Guatemala 1983 | <i>Afghanistan 1992</i> | Peru 2001 |
| <i>Iraq 1968</i> | Congo Brazzaville 1977 | Panama 1983 | Albania 1992 | Kenya 2002 |
| Mali 1968 | Spain 1977 | Swaziland 1983 | Algeria 1992 | Lesotho 2002 |
| Sierra Leone 1968 | Thailand 1977 | Thailand 1983 | <i>Azerbaijan 1992</i> | Sierra Leone 2002 |
| Benin 1969 | <i>Yemen North 1977</i> | Turkey 1983 | <i>Georgia 1992</i> | Ecuador 2003 |
| Bolivia 1969 | <i>Afghanistan 1978</i> | El Salvador 1984 | Guyana 1992 | Georgia 2003 |
| Brazil 1969 | Bolivia 1978 | Guinea 1984 | <i>Lithuania 1992</i> | Guinea-Bissau 2003 |
| <i>Libya 1969</i> | Comoros 1978 | Mauritania 1984 | Sierra Leone 1992 | <i>Iraq 2003</i> |
| <i>Yemen South 1969</i> | Ghana 1978 | Mongolia 1984 | <i>Tajikistan 1992</i> | Paraguay 2003 |
| Argentina 1970 | Honduras 1978 | Nicaragua 1984 | Thailand 1992 | Haiti 2004 |
| Bolivia 1970 | Mauritania 1978 | Nigeria 1985 | <i>Azerbaijan 1993</i> | |
| Ghana 1970 | <i>Yemen North 1978</i> | Sudan 1985 | Burundi 1993 | |

Note: Authoritarian breakdowns resulting in a transition to democracy are shown in **bold**. Authoritarian breakdowns shown in *italic* are omitted from our main regression (Model 1 of Table 5) due to missing data on explanatory variables.

Table A3: Effect of Diffusion on Democratization (Does Not Control for % Democracies in the World)

| | Spatial Lag 1 | | Spatial Lag 2 | | Spatial Lag 2 | | Spatial Lag 2 | |
|--------------------------------|--------------------|--------------------|--------------------|--------------------|--------------------|--------------------|--------------------|--------------------|
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| % Dem. neigh. | .532 (.147)*** | | | | .702 (.215)*** | | | |
| Ch. % dem. neigh. last year | | .607 (.338)* | | | | .810 (.579) | | |
| Ch. % dem. neigh. last 2 years | | | .359 (.251) | | | | .778 (.456)* | |
| Ch. % dem. neigh. last 3 years | | | | .161 (.220) | | | | .643 (.424) |
| Hegemonic power volatility | .207 (.077)*** | .190 (.078)** | .187 (.079)** | .198 (.079)** | .203 (.075)*** | .185 (.077)** | .176 (.079)** | .184 (.079)** |
| Growth rate | -.028 (.010)*** | -.028 (.010)*** | -.026 (.010)*** | -.029 (.010)*** | -.027 (.010)*** | -.028 (.010)*** | -.026 (.010)*** | -.028 (.010)*** |
| GDP pc (logged) | .243 (.071)*** | .315 (.067)*** | .316 (.067)*** | .317 (.067)*** | .217 (.081)*** | .316 (.067)*** | .317 (.067)*** | .315 (.067)*** |
| Oil | -.0003 (.0002)* | -.0004 (.0002)* | -.0004 (.0002)* | -.0004 (.0002)* | -.0003 (.0002) | -.0004 (.0002)* | -.0004 (.0002)* | -.0004 (.0002)* |
| Muslim | -.0009 (.002) | -.001 (.002) | -.0009 (.002) | -.001 (.002) | -.001 (.002) | -.001 (.002) | -.0009 (.002) | -.0009 (.002) |
| Catholic | -.0004 (.002) | -.0001 (.002) | -.0001 (.002) | -.0001 (.002) | -.0006 (.002) | -.0001 (.002) | -.00005 (.002) | -.00005 (.002) |
| Protestant | .003 (.002) | .003 (.002) | .003 (.002) | .003 (.002) | .004 (.002)* | .003 (.002) | .003 (.002) | .004 (.002) |
| Ethnic frac. | .003 (.002)* | .003 (.002)* | .003 (.002)* | .003 (.002)* | .003 (.002)* | .003 (.002)* | .003 (.002)* | .003 (.002)* |
| Religious frac. | -.003 (.003) | -.003 (.003) | -.003 (.003) | -.003 (.003) | -.003 (.003) | -.003 (.003) | -.003 (.003) | -.003 (.003) |
| Brit. col. | .057 (.130) | .069 (.130) | .073 (.131) | .075 (.130) | .109 (.130) | .071 (.130) | .073 (.131) | .072 (.131) |
| # Past trans. | .265 (.046)*** | .285 (.042)*** | .273 (.041)*** | .283 (.042)*** | .252 (.045)*** | .288 (.042)*** | .275 (.041)*** | .281 (.041)*** |
| N | 5621 | 5587 | 5526 | 5459 | 5621 | 5587 | 5526 | 5459 |
| Pseudolog-lik. | -475.964 | -477.259 | -474.608 | -476.088 | -476.164 | -477.63 | -474.253 | -475.321 |

Note: Redoes Table 4 without controlling for the % of democracies in the world. Dynamic probit estimations. Robust standard errors clustered by country in parentheses. All explanatory variables are lagged. *** $p < .01$, ** $p < .05$ and * $p < .1$.

Table A4: Effect of World Diffusion on Democratization

| | (1) | (2) | (3) | (4) |
|-------------------------------|--------------------|--------------------|--------------------|--------------------|
| % Dem. world | 1.622 (.397)*** | | | |
| Ch. % dem. world last year | | 2.703 (2.551) | | |
| Ch. % dem. world last 2 years | | | 4.111 (1.827)** | |
| Ch. % dem. world last 3 years | | | | 1.809 (1.427) |
| Hegemonic power volatility | .14 (.066)** | .142 (.072)** | .083 (.077) | .124 (.08) |
| Growth rate | -.029 (.010)*** | -.033 (.010)*** | -.031 (.010)*** | -.032 (.010)*** |
| GDP pc (logged) | .277 (.070)*** | .332 (.067)*** | .324 (.067)*** | .328 (.068)*** |
| Oil | -.0004 (.0002)* | -.0004 (.0002)* | -.0004 (.0002)* | -.0004 (.0002)* |
| Muslim | -.0009 (.002) | -.0002 (.002) | -.0003 (.002) | -.0002 (.002) |
| Catholic | .002 (.002) | .001 (.002) | .001 (.002) | .001 (.002) |
| Protestant | .003 (.003) | .002 (.003) | .002 (.003) | .002 (.003) |
| Ethnic frac. | .001 (.002) | .002 (.002) | .002 (.002) | .002 (.002) |
| Religion frac. | -.002 (.003) | -.002 (.003) | -.002 (.003) | -.002 (.003) |
| Brit. col. | -.012 (.141) | .049 (.137) | .038 (.137) | .045 (.137) |
| # Past trans. | .269 (.044)*** | .290 (.043)*** | .288 (.043)*** | .289 (.043)*** |
| N | 6477 | 6477 | 6477 | 6477 |
| Pseudolog-lik. | -497.993 | -505.622 | -503.008 | -505.099 |

Note: Redoes Table 4 using world diffusion. Dynamic probit estimations. Robust standard errors clustered by country in parentheses. All explanatory variables are lagged. *** $p < .01$, ** $p < .05$ and * $p < .1$.

Table A5: Effect of Regional Diffusion on Democratization

| | (1) | (2) | (3) | (4) |
|--------------------------------|--------------------|--------------------|--------------------|--------------------|
| % Dem. region | .546 (.198)*** | | | |
| Ch. % dem. region last year | | 2.015 (.686)*** | | |
| Ch. % dem. region last 2 years | | | .888 (.621) | |
| Ch. % dem. region last 3 years | | | | .715 (.559) |
| % Dem. world | 1.242 (.413)*** | 1.506 (.392)*** | 1.532 (.390)*** | 1.547 (.398)*** |
| Hegemonic power volatility | .147 (.066)** | .126 (.067)* | .129 (.066)* | .131 (.067)** |
| Growth rate | -.030 (.010)*** | -.029 (.010)*** | -.030 (.010)*** | -.030 (.010)*** |
| GDP pc (logged) | .239 (.072)*** | .279 (.069)*** | .279 (.070)*** | .280 (.070)*** |
| Oil | -.0004 (.0002)* | -.0004 (.0002)* | -.0004 (.0002)* | -.0004 (.0002)* |
| Muslim | -.0006 (.002) | -.0008 (.002) | -.0008 (.002) | -.0007 (.002) |
| Catholic | .0008 (.002) | .002 (.002) | .002 (.002) | .002 (.002) |
| Protestant | .002 (.003) | .003 (.003) | .003 (.003) | .003 (.003) |
| Ethnic frac. | .001 (.002) | .001 (.002) | .001 (.002) | .001 (.002) |
| Religious frac. | -.001 (.003) | -.002 (.003) | -.002 (.003) | -.002 (.003) |
| Brit. col. | -.018 (.145) | -.005 (.142) | -.008 (.143) | -.010 (.144) |
| # Past trans. | .229 (.050)*** | .270 (.044)*** | .270 (.044)*** | .267 (.045)*** |
| N | 6470 | 6429 | 6381 | 6313 |
| Pseudolog-lik. | -493.666 | -494.317 | -495.855 | -495.123 |

Note: Note: Redoes Table 4 using regional diffusion. Dynamic probit estimations. Robust standard errors clustered by country in parentheses. All explanatory variables are lagged. *** $p < .01$, ** $p < .05$ and * $p < .1$.

Table A6: Effect of Diffusion and Shocks on Authoritarian Breakdowns and Transitions to Democracy (Does Not Control for % Democracies in the World)

| | Spatial Lag 1 | | | | Spatial Lag 2 | | | |
|-----------------------------|-------------------------|----------------------|-------------------------|----------------------|---------------|-----|-----|-----|
| | (1) | (2) | (3) | (4) | (1) | (2) | (3) | (4) |
| % Dem. neigh. | .100 (.105) | .915 (.277)*** | .071 (.148) | 1.222 (.355)*** | | | | |
| Ch. % dem. neigh. last year | | | | | | | | |
| Hegemonic power volatility | .153 (.048)*** | .127 (.121) | .153 (.048)*** | .126 (.118) | | | | |
| Growth rate | -.019 (.006)*** | -.035 (.020)* | -.019 (.006)*** | -.030 (.020) | | | | |
| GDP pc (logged) | -.079 (.053) | .603 (.201)*** | -.074 (.051) | .587 (.201)*** | | | | |
| Oil | -6.88e-06 (1.00e-05) | -.0004 (.0002)*** | -8.05e-06 (1.00e-05) | -.0004 (.0002)*** | | | | |
| Muslim | -.0007 (.001) | -.001 (.003) | -.0007 (.001) | -.001 (.003) | | | | |
| Catholic | .002 (.001)* | -.004 (.003) | .002 (.001)* | -.004 (.003) | | | | |
| Protestant | -.0004 (.002) | .003 (.005) | -.0004 (.002) | .004 (.005) | | | | |
| Ethnic frac. | .0007 (.001) | .007 (.003)** | .0006 (.001) | .007 (.003)** | | | | |
| Religious frac. | -.003 (.002)* | .002 (.004) | -.003 (.002)* | .002 (.004) | | | | |
| Brit. col. | -.084 (.093) | .321 (.246) | -.078 (.093) | .371 (.257) | | | | |
| # Past trans. | .157 (.049)*** | .197 (.082)** | .158 (.049)*** | .165 (.083)** | | | | |
| N | 5488 | 423 | 5488 | 423 | | | | |
| Pseudolog-lik. | -1590.302 | -1567.283 | -1589.545 | -1569.415 | | | | |

Note: Redoes Table 5 without controlling for the % of democracies in the world. Heckman probit estimations. The first stage runs a probit estimation of the likelihood that an autocracy breaks down. The second stage runs a probit estimation of the likelihood that an autocracy that has just broken down transitions to democracy as opposed to another authoritarian regime. Models on authoritarian breakdown include the age of the regime, its square and its cube. Robust standard errors clustered by country in parentheses. All explanatory variables are lagged. *** $p < .01$, ** $p < .05$ and * $p < .1$.

Table A7: Effect of Diffusion and Shocks on Authoritarian Breakdowns and Transitions to Democracy (Change in % Democratic Neighbors in the Last 2 and 3 Years)

| | Spatial Lag 1 | | | | Spatial Lag 2 | | | |
|--------------------------------|-------------------------|----------------------|-------------------------|----------------------|-------------------------|-------------------------|--------------|-------------|
| | (1) | (2) | (3) | (4) | (1) | (2) | (3) | (4) |
| | Auth. Break. | Tr. to Dem. | Auth. Break. | Tr. to Dem. | Auth. Break. | Tr. to Dem. | Auth. Break. | Tr. to Dem. |
| Ch. % dem. neigh. last 2 years | -.103 (.194) | .337 (.501) | .112 (.375) | -.013 (.853) | .244 (.342) | | | |
| Ch. % dem. neigh. last 3 years | | | | | | | | |
| % Dem. world | .390 (.235)* | 3.239 (.811)*** | .374 (.236) | 3.279 (.819)*** | .369 (.235) | 3.342 (.826)*** | | |
| Hegemonic power volatility | .154 (.049)*** | .043 (.139) | .151 (.050)*** | .048 (.138) | .149 (.051)*** | .055 (.140) | | |
| Growth rate | -.020 (.006)*** | -.018 (.022) | -.019 (.006)*** | -.019 (.022) | -.019 (.006)*** | -.019 (.022) | | |
| GDP pc (logged) | -.080 (.052) | .832 (.222)*** | -.082 (.052) | .831 (.223)*** | -.082 (.052) | .849 (.220)*** | | |
| Oil | -7.32e-06 (1.00e-05) | -.0006 (.0002)*** | -5.90e-06 (1.00e-05) | -.0006 (.0002)*** | -7.26e-06 (1.00e-05) | -5.94e-06 (1.00e-05) | | |
| Muslim | -.0008 (.001)** | -.0008 (.004) | -.001 (.001)** | -.0004 (.004) | -.0008 (.001)** | -.001 (.001)** | | |
| Catholic | .003 (.001)** | -.002 (.004) | .003 (.001)** | -.002 (.004) | .003 (.001)** | -.002 (.004) | | |
| Protestant | -.0007 (.002) | .009 (.006) | -.0006 (.002) | .009 (.007) | -.0008 (.002) | .009 (.007) | | |
| Ethnic frac. | .0004 (.002) | .006 (.003)* | .0006 (.002) | .006 (.003)* | .0005 (.002) | .006 (.003)* | | |
| Religious frac. | -.004 (.002)* | .0004 (.005) | -.004 (.002)* | .0003 (.005) | -.004 (.002)* | .0003 (.005) | | |
| Brit. col. | -.082 (.093) | .243 (.288) | -.069 (.092) | .235 (.289) | -.082 (.093) | .239 (.288) | | |
| # Past trans. | .152 (.049)*** | .126 (.105) | .148 (.049)*** | .133 (.109) | .150 (.049)*** | .148 (.108) | | |
| N | 5402 | 415 | 5339 | 411 | 5402 | 415 | 5339 | 411 |
| Pseudolog-lik. | | -1550.763 | | -1534.901 | | -1550.988 | | -1534.568 |

Note: Redoes Table 5 using the change in % democracies in the last 2 and 3 years. Heckman probit estimations. The first stage runs a probit estimation of the likelihood that an autocracy breaks down. The second stage runs a probit estimation of the likelihood that an autocracy that has just broken down transitions to democracy as opposed to another authoritarian regime. Models on authoritarian breakdown include the age of the regime, its square and its cube. Robust standard errors clustered by country in parentheses. All explanatory variables are lagged. *** $p < .01$, ** $p < .05$ and * $p < .1$.

Table A8: Effect of Regional Diffusion and Shocks on Authoritarian Breakdowns and Transitions to Democracy

| | (1) | | (2) | | (3) | | (4) | |
|--------------------------------|-------------------------|----------------------|----------------------|----------------------|-------------------------|----------------------|-------------------------|----------------------|
| | Auth. Break. | Tr. to Dem. | Auth. Break. | Tr. to Dem. | Auth. Break. | Tr. to Dem. | Auth. Break. | Tr. to Dem. |
| % Dem. region | .101 (.141) | .734 (.333)** | | | | | | |
| Ch. % dem. region last year | | .814 (.492)* | 2.023 (1.473) | | | | | |
| Ch. % dem. region last 2 years | | | | 1.214 (.870) | | | | |
| Ch. % dem. region last 3 years | | | | | .360 (.494) | | | |
| % Dem. world | .438 (.244)* | 2.470 (.745)*** | 2.976 (.727)*** | 2.989 (.737)*** | .494 (.227)** | 2.989 (.737)*** | .212 (.407) | 1.193 (.954) |
| Hegemonic power volatility | .094 (.043)** | .086 (.112) | .059 (.116) | .053 (.119) | .093 (.044)** | .053 (.119) | .514 (.230)** | 3.004 (.736)*** |
| Growth rate | -.019 (.006)*** | -.027 (.020) | -.024 (.020) | -.025 (.021) | -.019 (.006)** | -.025 (.021) | .097 (.045)** | .051 (.121) |
| GDP pc (logged) | -.066 (.051) | .645 (.189)*** | .725 (.187)*** | .749 (.190)*** | -.062 (.050) | .749 (.190)*** | -.059 (.051) | .762 (.191)*** |
| Oil | -6.18e-06 (1.00e-05) | -.0005 (.0002)*** | -.0006 (.0002)*** | -.0006 (.0002)*** | -6.01e-06 (1.00e-05) | -.0006 (.0002)*** | -6.04e-06 (1.00e-05) | -.0006 (.0002)*** |
| Muslim | -.0001 (.001) | -.0009 (.003) | -.0008 (.003) | -.0008 (.003) | -.0002 (.001) | -.0004 (.003) | -.0002 (.001) | -.0003 (.003) |
| Catholic | .004 (.001)*** | -.003 (.003) | -.001 (.003) | -.001 (.003) | .004 (.001)*** | -.001 (.003) | .004 (.001)*** | -.001 (.003) |
| Protestant | -.002 (.002) | .002 (.005) | .004 (.006) | .005 (.006) | -.001 (.002) | .005 (.006) | -.001 (.002) | .005 (.006) |
| Ethnic frac. | -.0003 (.002) | .004 (.003) | .004 (.003) | .005 (.003) | -.0003 (.001) | .005 (.003) | -.0003 (.002) | .005 (.003) |
| Religious frac. | -.002 (.002) | .001 (.004) | .003 (.004) | .003 (.004) | -.002 (.002) | .005 (.004) | -.002 (.002) | .001 (.004) |
| Brit. col. | -.131 (.101) | .175 (.279) | .190 (.281) | .180 (.287) | -.124 (.101) | .190 (.287) | -.108 (.102) | .171 (.291) |
| # Past trans. | .139 (.048)*** | .148 (.087)* | .175 (.091)* | .175 (.098)* | .145 (.047)*** | .175 (.098)* | .144 (.048)*** | .170 (.098)* |
| N | 6200 | 450 | 450 | 447 | 6126 | 447 | 6064 | 443 |
| Pseudolog-lik. | | -1701.985 | -1699.506 | -1689.096 | | -1673.735 | | |

Note: Redoes Table 5 using regional diffusion. Heckman probit estimations. The first stage runs a probit estimation of the likelihood that an autocracy breaks down. The second stage runs a probit estimation of the likelihood that an autocracy that has just broken down transitions to democracy as opposed to another authoritarian regime. Models on authoritarian breakdown include the age of the regime, its square and its cube. Robust standard errors clustered by country in parentheses. All explanatory variables are lagged. *** $p < .01$, ** $p < .05$ and * $p < .1$.

Table A9: Effect of Diffusion and Shocks on Authoritarian Breakdowns and Transitions to Democracy (Change % World Democracies in the Last Year)

| | Spatial Lag 1 | | Spatial Lag 2 | | | |
|-----------------------------|--------------------------------|----------------------------------|--------------------------------|-----------------------------------|----------------------------------|----------------------------------|
| | (1) | (2) | (3) | (4) | | |
| % Dem. neigh. | Auth. Break. .099 (.104) | Tr. to Dem. .924 (.283)*** | Auth. Break. .071 (.147) | Tr. to Dem. 1.232 (.365)*** | Auth. Break. -0.423 (.525) | Tr. to Dem. -0.091 (1.118) |
| Ch. % dem. neigh. last year | | | | | | |
| Ch. % dem. world last year | 1.220 (1.688) | 3.339 (3.886) | 1.241 (1.690) | 3.224 (3.968) | 2.027 (1.771) | 2.712 (4.011) |
| Hegemonic power volatility | .140 (.053)*** | .090 (.132) | .140 (.054)*** | .090 (.130) | .146 (.053)*** | .120 (.132) |
| Growth rate | -.019 (.006)*** | -.034 (.021) | -.019 (.006)*** | -.028 (.021) | -.019 (.006)*** | -.029 (.021) |
| GDP pc (logged) | -.080 (.053) | .612 (.202)*** | -.070 (.051) | .594 (.203)*** | -.070 (.051) | .843 (.218)*** |
| Oil | -7.00e-06 (1.00e-05) | -.0004 (.0002)*** | -8.14e-06 (1.00e-05) | -.0004 (.0002)*** | -8.88e-06 (1.00e-05) | -.0006 (.0002)*** |
| Muslim | -.0007 (.001) | -.001 (.003) | -.0007 (.001) | -.001 (.003) | -.0006 (.001) | -.0006 (.004) |
| Catholic | .002 (.001)* | -.004 (.003) | .002 (.001)* | -.004 (.003) | .003 (.001)** | -.004 (.003) |
| Protestant | -.0005 (.002) | .003 (.005) | -.0004 (.002) | .006 (.006) | -.001 (.002) | .006 (.006) |
| Ethnic frac. | .0006 (.001) | .007 (.003)** | .0006 (.002) | .007 (.003)** | .0006 (.002) | .007 (.003)** |
| Religious frac. | -.003 (.002)* | .002 (.004) | -.003 (.002)* | .002 (.004) | -.003 (.002)* | .002 (.005) |
| Brit. col. | -.086 (.093) | .324 (.250) | -.081 (.093) | .374 (.260) | -.083 (.093) | .331 (.248) |
| # Past trans. | .156 (.049)*** | .193 (.084)** | .157 (.049)*** | .162 (.085)* | .162 (.049)*** | .177 (.100)* |
| N | 5488 | 423 | 5488 | 423 | 5458 | 416 |
| Pseudolog-lik. | | -1589.806 | | -1589.064 | | -1568.663 |

Note: Redoes Table 5 using *Change % World Democracies in the last year* rather than *% World Democracies*. Heckman probit estimations. The first stage runs a probit estimation of the likelihood that an autocracy breaks down. The second stage runs a probit estimation of the likelihood that an autocracy that has just broken down transitions to democracy as opposed to another authoritarian regime. Models on authoritarian breakdown include the age of the regime, its square and its cube. Robust standard errors clustered by country in parentheses. All explanatory variables are lagged. *** $p < .01$, ** $p < .05$ and * $p < .1$.

Table A10: Effect of Alternative Economic Shocks on Authoritarian Breakdowns and Transitions to Democracy (Does Not Control for % Democracies in the World)

| | (1) | (2) | (3) | (4) |
|-----------------------------|-------------------------|---------------------|--------------------|--------------------|
| | Auth. Break. | Auth. Break. | Auth. Break. | Auth. Break. |
| | Tr. to Dem. | Tr. to Dem. | Tr. to Dem. | Tr. to Dem. |
| % Dem. neigh. | .054 (.114) | .950 (.350)*** | .064 (.120) | .879 (.279)*** |
| Ch. % dem. neigh. last year | | 1.880 (.690)*** | | -114 (.242) |
| Hegemonic power volatility | .264 (.083)*** | .515 (.233)** | .161 (.055)*** | .175 (.055)*** |
| Rain dev. | -.049 (.015)*** | -.035 (.050) | | |
| RR/LV crises | | | .226 (.056)*** | .231 (.056)*** |
| GDP pc (logged) | -.109 (.051)** | .512 (.242)** | -.071 (.060) | -.067 (.056) |
| Oil | -7.09e-06 (1.00e-05) | -.0004 (.0002)** | -.0002 (.00003) | -.0002 (.00003) |
| Muslim | .0005 (.002) | -.0004 (.003) | -.0009 (.002) | -.0009 (.002) |
| Catholic | .004 (.002)** | -.0008 (.004) | .002 (.001) | .002 (.001) |
| Protestant | -.002 (.003) | -.003 (.006) | -.001 (.002) | -.002 (.002) |
| Ethnic frac. | .00006 (.002) | .010 (.004)** | .0009 (.002) | .001 (.002) |
| Religious frac. | -.0001 (.002) | .004 (.005) | -.003 (.002) | -.003 (.002) |
| Brit. col. | -.101 (.091) | .388 (.296) | -.070 (.100) | -.072 (.100) |
| # Past trans. | .173 (.057)*** | .166 (.121) | .163 (.052)*** | .166 (.052)*** |
| N | 3384 | 250 | 4685 | 4664 |
| Pseudolog-lik. | -973.717 | -959.065 | -1409.389 | -1394.649 |

Note: Redoes Table 6 without controlling for the % of democracies in the world. Heckman probit estimations. The first stage runs a probit estimation of the likelihood that an autocracy breaks down. The second stage runs a probit estimation of the likelihood that an autocracy that has just broken down transitions to democracy as opposed to another authoritarian regime. Uses spatial lag 1. *Rain deviation* gives the difference, in percentage, between the total amount of rain a country received in a given year and the average yearly amount of rain that country has received between 1950 and 2006. The variable *RR/LV crises* indicates whether a country is experiencing a crisis as defined by either Rogoff and Reinhart (2010) or Laeven and Valencia (2013). Models on authoritarian breakdown include the age of the regime, its square and its cube. Robust standard errors clustered by country in parentheses. All explanatory variables are lagged. *** $p < .01$, ** $p < .05$ and * $p < .1$.

Table A11: Effect of Alternative Economic Shocks on Authoritarian Breakdowns and Transitions to Democracy (RR and LV Crises)

| | (1) | (2) | (3) | (4) |
|-----------------------------|----------------------|--------------------|-------------------|--------------------|
| | Auth. Break. | Auth. Break. | Auth. Break. | Auth. Break. |
| | Tr. to Dem. | Tr. to Dem. | Tr. to Dem. | Tr. to Dem. |
| % Dem. neigh. | -.087 (.150) | .458 (.366) | .022 (.152) | .750 (.433)* |
| Ch. % dem. neigh. last year | | | | |
| % Dem. world | .399 (.364) | 2.882 (1.162)** | -.302 (.507) | 3.206 (1.311)** |
| Hegemonic power volatility | .125 (.066)* | -.160 (.170) | .233 (.115)** | .699 (.313)** |
| RR crises | .154 (.068)** | .054 (.214) | .153 (.069)** | .019 (.211) |
| LV crises | | | | |
| GDP pc (logged) | .041 (.077) | .787 (.219)*** | .293 (.090)*** | -.030 (.359) |
| Oil | -2.66e-06 (.0002) | -0.003 (.0002)* | -.137 (.067)** | .727 (.251)*** |
| Muslim | -.003 (.002) | -.005 (.006) | -.002 (.002) | -.002 (.002) |
| Catholic | .003 (.002) | -.004 (.003) | .003 (.003) | .003 (.003) |
| Protestant | -.006 (.003) | .016 (.008)* | -.006 (.003)* | -.012 (.011) |
| Ethnic frac. | .004 (.002)* | .007 (.005) | -.002 (.002) | .013 (.005)** |
| Religious frac. | -.009 (.003)*** | .002 (.008) | -.005 (.003) | -.003 (.007) |
| Brit. col. | -.185 (.169) | .549 (.453) | -.061 (.105) | -.070 (.106) |
| # Past trans. | .171 (.056)*** | -.012 (.155) | .231 (.067)*** | .053 (.210) |
| N | 3241 | 284 | 2326 | 166 |
| Pseudolog-lik. | | -1023.812 | -639.966 | -633.811 |

Note: Redoes Table 6 with *RR crises* and *LV crises* on their own. Heckman probit estimations. The first stage runs a probit estimation of the likelihood that an autocracy breaks down. The second stage runs a probit estimation of the likelihood that an autocracy that has just broken down transitions to democracy as opposed to another authoritarian regime. Uses spatial lag 1. *Rain deviation* gives the difference, in percentage, between the total amount of rain a country received in a given year and the average yearly amount of rain that country has received between 1950 and 2006. The variable *RR/LV crises* indicates whether a country is experiencing a crisis as defined by either Rogoff and Reinhart (2010) or Laeven and Valencia (2013). Models on authoritarian breakdown include the age of the regime, its square and its cube. Robust standard errors clustered by country in parentheses. All explanatory variables are lagged. *** $p < .01$, ** $p < .05$ and * $p < .1$.

Table A12: Effect of Alternative Economic Shocks on Authoritarian Breakdowns and Transitions to Democracy (Change % Democracies over 2 and 3 Years)

| | (1) | | (2) | | (3) | | (4) | |
|--------------------------------|-------------------------|----------------------|-------------------------|----------------------|---------------------|---------------------|---------------------|---------------------|
| | Auth. Break. | Tr. to Dem. | Auth. Break. | Tr. to Dem. | Auth. Break. | Tr. to Dem. | Auth. Break. | Tr. to Dem. |
| Ch. % dem. neigh. last 2 years | -.012 (.281) | .660 (.663) | | | .056 (.199) | .229 (.521) | | |
| Ch. % dem. neigh. last 3 years | | | .100 (.286) | -.152 (.522) | | | .141 (.223) | -.354 (.406) |
| % Dem. world | -.173 (.399) | 4.273 (.967)*** | -.141 (.405) | 4.398 (1.004)*** | .234 (.263) | 3.203 (.866)*** | .254 (.266) | 3.281 (.849)*** |
| Hegemonic power volatility | .276 (.081)*** | .309 (.286) | .294 (.082)*** | .278 (.284) | .169 (.055)*** | .049 (.151) | .173 (.057)*** | .050 (.152) |
| Rain dev. | -.050 (.015)*** | .040 (.055) | -.055 (.015)*** | .052 (.056) | | | | |
| RR/LV crises | | | | | .220 (.059)*** | -.070 (.195) | .207 (.059)*** | -.062 (.191) |
| GDP pc (logged) | -.117 (.049)** | .836 (.167)*** | -.115 (.049)** | .847 (.164)*** | -.074 (.057) | .854 (.223)*** | -.075 (.057) | .877 (.219)*** |
| Oil | -6.69e-06 (1.00e-05) | -.0006 (.0002)*** | -5.82e-06 (1.00e-05) | -.0006 (.0002)*** | -.00002 (.00003) | -.0006 (.0002)** | -.00002 (.00003) | -.0006 (.0002)** |
| Muslim | .0004 (.002) | -.0003 (.004) | .0003 (.002) | -.0005 (.004) | -.001 (.002) | -.001 (.004) | -.001 (.002) | -.0006 (.004) |
| Catholic | .005 (.002)** | -.002 (.004) | .005 (.002)** | -.003 (.004) | .002 (.001) | -.004 (.003) | .002 (.002) | -.004 (.003) |
| Protestant | -.004 (.003) | .0001 (.009) | -.004 (.003) | .001 (.009) | -.002 (.002) | .008 (.007) | -.002 (.002) | .008 (.007) |
| Ethnic frac. | .0005 (.002) | .008 (.005) | .007 (.002) | .008 (.005) | .0008 (.002) | .007 (.004)* | .001 (.002) | .007 (.004)* |
| Religious | .0002 (.002) | -.0001 (.006) | .0003 (.002) | -.0003 (.006) | -.003 (.002) | .003 (.005) | -.003 (.002) | .003 (.005) |
| Brit. col. | -.080 (.088) | .401 (.338) | -.064 (.088) | .369 (.338) | -.078 (.099) | .185 (.309) | -.070 (.100) | .161 (.312) |
| # Past trans. | .164 (.056)*** | .033 (.169) | .157 (.056)*** | .047 (.167) | .156 (.051)*** | .110 (.111) | .155 (.051)*** | .115 (.118) |
| N | 3311 | 244 | 3251 | 240 | 4630 | 371 | 4593 | 368 |
| Pseudolog-lik. | | -940.406 | | -923.869 | | -1384.936 | | -1373.239 |

Note: Redoes Table 6 with change % democracies over 2 and 3 Years. Heckman probit estimations. The first stage runs a probit estimation of the likelihood that an autocracy breaks down. The second stage runs a probit estimation of the likelihood that an autocracy that has just broken down transitions to democracy as opposed to another authoritarian regime. Uses spatial lag 1. *Rain deviation* gives the difference, in percentage, between the total amount of rain a country received in a given year and the average yearly amount of rain that country has received between 1950 and 2006. The variable *RR/LV crises* indicates whether a country is experiencing a crisis as defined by either Rogoff and Reinhart (2010) or Laeven and Valencia (2013). Models on authoritarian breakdown include the age of the regime, its square and its cube. Robust standard errors clustered by country in parentheses. All explanatory variables are lagged. *** $p < .01$, ** $p < .05$ and * $p < .1$.

Table A13: Effect of Alternative Economic Shocks on Authoritarian Breakdowns and Transitions to Democracy (Spatial Lag 2)

| | (1) | (2) | (3) | (4) |
|-----------------------------|-------------------------|-------------------------|---------------------|---------------------|
| | Auth. Break. | Auth. Break. | Auth. Break. | Auth. Break. |
| | Tr. to Dem. | Tr. to Dem. | Tr. to Dem. | Tr. to Dem. |
| | Auth. Break. | Auth. Break. | Auth. Break. | Auth. Break. |
| | Tr. to Dem. | Tr. to Dem. | Tr. to Dem. | Tr. to Dem. |
| % Dem. neigh. | .056 (.192) | .576 (.542) | .065 (.183) | .874 (.388)** |
| Ch. % dem. neigh. last year | | .540 (.588) | | -.143 (.506) |
| % Dem. world | -.201 (.399) | 3.605 (1.133)** | .173 (.273) | 2.098 (.874)** |
| Hegemonic power volatility | .269 (.082)** | .265 (.080)** | .154 (.054)** | .166 (.055)** |
| Rain dev. | -.051 (.015)** | -.053 (.015)** | | |
| RR/LV crises | | | .221 (.057)** | -.007 (.169) |
| GDP pc (logged) | -.108 (.053)** | -.118 (.049)** | -.074 (.061) | -.074 (.057) |
| Oil | -7.75e-06 (1.00e-05) | -6.65e-06 (1.00e-05) | -.00002 (.00003) | -.00002 (.00003) |
| Muslim | .0005 (.002) | -.0005 (.004) | -.001 (.002) | -.001 (.002) |
| Catholic | .004 (.002)** | .005 (.002)** | .002 (.001) | .002 (.001) |
| Protestant | -.002 (.003) | -.004 (.003) | -.001 (.002) | -.002 (.002) |
| Ethnic frac. | .00009 (.002) | .009 (.005)* | .008 (.002) | .009 (.002) |
| Religious frac. | -.00006 (.002) | .002 (.002) | -.003 (.002) | -.003 (.002) |
| Brit. col. | -.096 (.092) | -.085 (.088) | -.071 (.101) | -.079 (.100) |
| # Past trans. | .174 (.056)** | .170 (.056)** | .158 (.051)** | .161 (.051)** |
| N | 3384 | 250 | 4685 | 375 |
| Pseudolog-lik. | -968.456 | -951.31 | -1405.013 | -1388.311 |

Note: Redoes Table 6 with spatial lag 2. Heckman probit estimations. The first stage runs a probit estimation of the likelihood that an autocracy breaks down. The second stage runs a probit estimation of the likelihood that an autocracy that has just broken down transitions to democracy as opposed to another authoritarian regime. *Rain deviation* gives the difference, in percentage, between the total amount of rain a country received in a given year and the average yearly amount of rain that country has received between 1950 and 2006. The variable *RR/LV crises* indicates whether a country is experiencing a crisis as defined by either Rogoff and Reinhart (2010) or Laeven and Valencia (2013). Models on authoritarian breakdown include the age of the regime, its square and its cube. Robust standard errors clustered by country in parentheses. All explanatory variables are lagged. *** $p < .01$, ** $p < .05$ and * $p < .1$. *** $p < .01$, ** $p < .05$ and * $p < .1$.

Table A14: Effect of Alternative Economic Shocks on Authoritarian Breakdowns and Transitions to Democracy (Regional Diffusion)

| | (1) | (2) | (3) | (4) |
|-----------------------------|-------------------------|----------------------|--------------------|--------------------|
| | Auth. Break. | Tr. to Dem. | Auth. Break. | Tr. to Dem. |
| | Auth. Break. | Tr. to Dem. | Auth. Break. | Tr. to Dem. |
| % Dem. region | .161 (.177) | .299 (.512) | .083 (.165) | .728 (.350)** |
| Ch. % dem. region last year | | | | |
| % Dem. world | -.291 (.399) | 3.948 (1.035)*** | .452 (.647) | 2.154 (1.772) |
| Hegemonic power volatility | .274 (.082)*** | .412 (.274) | .172 (.274) | 2.413 (.814)*** |
| Rain dev. | -.049 (.015)*** | .021 (.055) | .14 (.051)*** | .131 (.051)*** |
| RR/LV crises | | | | |
| GDP pc (logged) | -.113 (.048)** | .727 (.186)*** | .226 (.055)*** | .227 (.055)*** |
| Oil | -6.16e-06 (1.00e-05) | -.0006 (.0002)*** | -.073 (.059) | -.066 (.056) |
| Muslim | .0004 (.002) | -.002 (.004) | -.0007 (.001) | -.0006 (.001) |
| Catholic | .004 (.002)** | -.002 (.004) | .003 (.0009)*** | .003 (.003) |
| Protestant | -.003 (.003) | -.006 (.008) | -.002 (.002) | -.002 (.002) |
| Ethnic frac. | -.00004 (.002) | .008 (.005) | .0005 (.002) | .0004 (.002) |
| Religious frac. | .0006 (.002) | -.0002 (.006) | -.002 (.002) | -.002 (.002) |
| Brit. col. | -.124 (.089) | .387 (.343) | -.078 (.094) | -.068 (.092) |
| # Past trans. | .162 (.056)*** | .091 (.167) | .149 (.050)*** | .153 (.049)*** |
| N | 3476 | 255 | 4993 | 4974 |
| Pseudolog-lik. | -988.2 | -977.878 | -1485.904 | -1481.834 |

Note: Redoes Table 6 with regional diffusion. Heckman probit estimations. The first stage runs a probit estimation of the likelihood that an autocracy breaks down. The second stage runs a probit estimation of the likelihood that an autocracy that has just broken down transitions to democracy as opposed to another authoritarian regime. *Rain deviation* gives the difference, in percentage, between the total amount of rain a country received in a given year and the average yearly amount of rain that country has received between 1950 and 2006. The variable *RR/LV crises* indicates whether a country is experiencing a crisis as defined by either Rogoff and Reinhart (2010) or Laeven and Valencia (2013). Models on authoritarian breakdown include the age of the regime, its square and its cube. Robust standard errors clustered by country in parentheses. All explanatory variables are lagged. *** $p < .01$, ** $p < .05$ and * $p < .1$. *** $p < .01$, ** $p < .05$ and * $p < .1$.

Table A15: Effect of Alternative Economic Shocks on Authoritarian Breakdowns and Transitions to Democracy (Change % World Democracies in the Last Year)

| | (1) | (2) | (3) | (4) |
|-----------------------------|-------------------------|-------------------------|---------------------|---------------------|
| | Auth. Break. | Auth. Break. | Auth. Break. | Auth. Break. |
| | Tr. to Dem. | Tr. to Dem. | Tr. to Dem. | Tr. to Dem. |
| % Dem. neigh. | .055 (.113) | 1.016 (.356)*** | .063 (.119) | .882 (.282)*** |
| Ch. % dem. neigh. last year | | .086 (.346) | 1.759 (.676)*** | -1.173 (.256) |
| Ch. % dem. world last year | .565 (2.310) | 1.167 (2.362) | 1.484 (1.794) | 1.847 (1.849) |
| Hegemonic power volatility | .257 (.089)*** | .252 (.087)*** | .144 (.061)** | .154 (.061)** |
| Rain dev. | -.048 (.015)*** | -.050 (.016)*** | -.015 (.049) | .141 (.145) |
| RR/LV crises | | | .225 (.057)*** | .230 (.057)*** |
| GDP pc (logged) | -.110 (.051)** | -.119 (.049)** | -.072 (.059) | -.067 (.055) |
| Oil | -7.09e-06 (1.00e-05) | -6.55e-06 (1.00e-05) | -.00002 (.00003) | -.00002 (.00003) |
| Muslim | .0005 (.002) | -.001 (.004) | -.0009 (.002) | -.0009 (.002) |
| Catholic | .004 (.002)** | -.002 (.004) | .002 (.001) | .002 (.001) |
| Protestant | -.002 (.003) | -.004 (.007) | -.001 (.002) | -.002 (.002) |
| Ethnic frac. | .00004 (.002) | .010 (.004)** | .008 (.002) | .009 (.002) |
| Religious frac. | -.0001 (.002) | .003 (.006) | -.003 (.002) | -.003 (.002) |
| Brit. col. | -.101 (.091) | .416 (.301) | -.072 (.100) | -.074 (.100) |
| # Past trans. | .172 (.057)*** | .136 (.139) | .161 (.052)*** | .165 (.052)*** |
| N | 3384 | 250 | 4685 | 4664 |
| Pseudolog-lik. | -971.819 | -957.933 | -1409.039 | -1394.11 |

Note: Redoes Table 5 using *Change % World Democracies in the last year* rather than *% World Democracies*. Heckman probit estimations. The first stage runs a probit estimation of the likelihood that an autocracy breaks down. The second stage runs a probit estimation of the likelihood that an autocracy that has just broken down transitions to democracy as opposed to another authoritarian regime. Uses spatial lag 1. *Rain deviation* gives the difference, in percentage, between the total amount of rain a country received in a given year and the average yearly amount of rain that country has received between 1950 and 2006. The variable *RR/LV crises* indicates whether a country is experiencing a crisis as defined by either Rogoff and Reinhart (2010) or Laeven and Valencia (2013). Models on authoritarian breakdown include the age of the regime, its square and its cube. Robust standard errors clustered by country in parentheses. All explanatory variables are lagged. *** $p < .01$, ** $p < .05$ and * $p < .1$.

Table A16: Effect of Diffusion and Shocks on Authoritarian Breakdowns and Transitions to Democracy (1900-2004)

| | Spatial Lag 1 | | | | Spatial Lag 2 | | | |
|-----------------------------|-------------------------|----------------------|-------------------------|----------------------|-------------------------|----------------------|-------------------------|----------------------|
| | (1) | (2) | (3) | (4) | (1) | (2) | (3) | (4) |
| | Auth. Break. | Tr. to Dem. | Auth. Break. | Tr. to Dem. | Auth. Break. | Tr. to Dem. | Auth. Break. | Tr. to Dem. |
| % Dem. neigh. | .098 (.107) | .576 (.288)** | | | .054 (.168) | .763 (.398)* | | |
| Ch. % dem. neigh. last year | | | -239 (.249) | 1.731 (.632)*** | | | -149 (.497) | -233 (1.021) |
| % Dem. world | .073 (.322) | 2.427 (.942)*** | .141 (.318) | 3.230 (1.005)*** | .100 (.327) | 2.249 (.976)** | .137 (.320) | 3.234 (1.000)*** |
| Hegemonic power volatility | .132 (.048)*** | .054 (.129) | .136 (.050)*** | .036 (.135) | .131 (.048)*** | .060 (.126) | .136 (.050)*** | .048 (.131) |
| Growth rate | -.019 (.006)*** | -.022 (.020) | -.019 (.006)*** | -.015 (.022) | -.019 (.006)*** | -.020 (.020) | -.019 (.006)*** | -.018 (.021) |
| GDP pc (logged) | -.082 (.053) | .574 (.214)*** | -.079 (.053) | .764 (.231)*** | -.077 (.055) | .565 (.217)*** | -.079 (.053) | .752 (.237)*** |
| Oil | -7.21e-06 (1.00e-05) | -.0005 (.0002)*** | -8.39e-06 (1.00e-05) | -.0006 (.0002)*** | -8.55e-06 (1.00e-05) | -.0005 (.0002)*** | -8.33e-06 (1.00e-05) | -.0006 (.0002)*** |
| Muslim | -.0009 (.001) | -.001 (.003) | -.0009 (.001) | -.001 (.004) | -.001 (.001) | -.001 (.003) | -.0009 (.001) | -.0008 (.004) |
| Catholic | .002 (.001) | -.002 (.003) | .002 (.001)* | -.002 (.004) | -.002 (.001) | -.002 (.003) | .002 (.001)* | -.002 (.004) |
| Protestant | .0007 (.003) | .003 (.006) | .0002 (.003) | .007 (.007) | .0008 (.003) | .004 (.006) | 2.19e-06 (.003) | .007 (.007) |
| Ethnic frac. | -.0002 (.001) | .005 (.003) | -.0003 (.002) | .005 (.004) | -.0003 (.002) | .005 (.003) | -.0003 (.002) | .005 (.003) |
| Religious frac. | -.004 (.002)** | -.0002 (.005) | -.004 (.002)** | .0004 (.005) | -.004 (.002)** | -.0001 (.005) | -.004 (.002)** | -.0004 (.005) |
| Brit. col. | -.107 (.094) | .237 (.279) | -.098 (.095) | .263 (.295) | -.103 (.095) | .266 (.286) | -.099 (.095) | .242 (.289) |
| # Past trans. | .148 (.049)*** | .188 (.088)** | .152 (.049)*** | .154 (.107) | .149 (.049)*** | .171 (.087)** | .152 (.049)*** | .162 (.104) |
| N | 4769 | 378 | 4744 | 373 | 4769 | 378 | 4744 | 373 |
| Pseudolog-lik. | | -1419.533 | | -1397.89 | | -1419.611 | | -1399.788 |

Note: Redoes Table 5 while restricting the sample between 1900 and 2004 (rather than 1875 and 2004). Heckman probit estimations. The first stage runs a probit estimation of the likelihood that an autocracy breaks down. The second stage runs a probit estimation of the likelihood that an autocracy that has just broken down transitions to democracy as opposed to another authoritarian regime. Models on authoritarian breakdown include the age of the regime, its square and its cube. Robust standard errors clustered by country in parentheses. All explanatory variables are lagged. *** $p < .01$, ** $p < .05$ and * $p < .1$.

Table A17: Effect of Diffusion and Shocks on Authoritarian Breakdowns and Transitions to Democracy (1900-2004; Does Not Control for % Democracies in the World)

| | Spatial Lag 1 | | | | Spatial Lag 2 | | | |
|-----------------------------|-------------------------|---------------------|-------------------------|----------------------|-------------------------|---------------------|-------------------------|----------------------|
| | (1) | (2) | (3) | (4) | (1) | (2) | (3) | (4) |
| | Auth. Break. | Tr. to Dem. | Auth. Break. | Tr. to Dem. | Auth. Break. | Tr. to Dem. | Auth. Break. | Tr. to Dem. |
| % Dem. neigh. | .105 (.103) | .806 (.271)*** | | | .071 (.158) | 1.099 (.364)*** | | |
| Ch. % dem. neigh. last year | | | -.231 (.247) | 1.728 (.614)*** | | | -.130 (.493) | .199 (.997) |
| Hegemonic power volatility | .133 (.049)*** | .096 (.118) | .138 (.050)*** | .098 (.119) | .133 (.049)*** | .098 (.116) | .138 (.050)*** | .104 (.116) |
| Growth rate | -.019 (.006)*** | -.032 (.019)* | -.020 (.006)*** | -.026 (.020) | -.019 (.006)*** | -.027 (.019) | -.020 (.006)*** | -.028 (.020) |
| GDP pc (logged) | -.081 (.053) | .512 (.201)** | -.075 (.053) | .717 (.215)*** | -.077 (.055) | .502 (.200)** | -.075 (.052) | .704 (.217)*** |
| Oil | -7.28e-06 (1.00e-05) | -.0004 (.0002)** | -8.80e-06 (1.00e-05) | -.0005 (.0002)*** | -8.55e-06 (1.00e-05) | -.0004 (.0002)** | -8.72e-06 (1.00e-05) | -.0005 (.0002)*** |
| Muslim | -.0009 (.001) | -.001 (.003) | -.0008 (.001) | -.0008 (.004) | -.001 (.001) | -.001 (.003) | -.0009 (.001) | -.0007 (.003) |
| Catholic | .002 (.001) | -.003 (.003) | .002 (.001)* | -.003 (.003) | .002 (.001) | -.003 (.003) | .002 (.001)* | -.002 (.003) |
| Protestant | .0007 (.003) | .002 (.005) | .0001 (.003) | .005 (.006) | .0008 (.003) | .003 (.005) | -3.51e-06 (.003) | .005 (.006) |
| Ethnic frac. | -.0002 (.001) | .006 (.003)* | -.0003 (.002) | .006 (.003)* | -.0003 (.001) | .006 (.003)* | -.0003 (.002) | .006 (.003)* |
| Religious frac. | -.004 (.002)** | .002 (.004) | -.004 (.002)** | .003 (.005) | -.004 (.002)** | .002 (.004) | -.004 (.002)** | .002 (.005) |
| Brit. col. | -.105 (.095) | .277 (.241) | -.095 (.096) | .292 (.242) | -.100 (.096) | .323 (.251) | -.096 (.096) | .277 (.238) |
| # Past trans. | .148 (.049)*** | .201 (.076)*** | .154 (.049)*** | .180 (.090)** | .149 (.049)*** | .173 (.076)** | .153 (.049)*** | .183 (.089)** |
| N | 4769 | 378 | 4744 | 373 | 4769 | 378 | 4744 | 373 |
| Pseudolog-lik. | -1423.469 | -1404.862 | -1422.78 | -1407.018 | -1422.78 | -1407.018 | -1422.78 | -1407.018 |

Note: Redoes Table 5 while restricting the sample between 1900 and 2004 (rather than 1875 and 2004), and without controlling for the % of democracies in the world. Heckman probit estimations. The first stage runs a probit estimation of the likelihood that an autocracy breaks down. The second stage runs a probit estimation of the likelihood that an autocracy that has just broken down transitions to democracy as opposed to another authoritarian regime. Models on authoritarian breakdown include the age of the regime, its square and its cube. Robust standard errors clustered by country in parentheses. All explanatory variables are lagged. *** $p < .01$, ** $p < .05$ and * $p < .1$.

Table A18: Effect of Diffusion and Shocks on Authoritarian Breakdowns and Transitions to Democracy (1945-2004)

| | Spatial Lag 1 | | Spatial Lag 2 | |
|-----------------------------|-------------------------|----------------------|-------------------------|----------------------|
| | (1) | (2) | (3) | (4) |
| % Dem. neigh. | .077 (.114) | .507 (.350) | -.020 (.190) | .695 (.535) |
| Ch. % dem. neigh. last year | | | | |
| % Dem. world | -.077 (.399) | 3.720 (1.010)*** | -.018 (.397) | 3.578 (1.094)*** |
| Hegemonic power volatility | .212 (.064)*** | .130 (.187) | .212 (.064)*** | .152 (.188) |
| Growth rate | -.016 (.007)** | -.008 (.026) | -.016 (.007)** | -.006 (.026) |
| GDP pc (logged) | -.099 (.050)** | .665 (.187)*** | -.092 (.052)* | .661 (.193)*** |
| Oil | -5.19e-06 (1.00e-05) | -.0006 (.0002)*** | -6.85e-06 (1.00e-05) | -.0006 (.0002)*** |
| Muslim | -.0001 (.002) | -.0008 (.004) | -.0002 (.002) | -.0007 (.004) |
| Catholic | .004 (.002)** | -.002 (.004) | .004 (.002)** | -.002 (.004) |
| Protestant | -.001 (.003) | -.004 (.008) | -.003 (.003) | -.003 (.003) |
| Ethnic frac. | -.0002 (.002) | .010 (.005)* | -.0003 (.002) | .010 (.005)* |
| Religious frac. | -.001 (.002) | -.0006 (.006) | -.002 (.002) | -.0007 (.006) |
| Brit. col. | -.103 (.092) | .393 (.333) | -.104 (.093) | .420 (.346) |
| # Past trans. | .153 (.052)*** | .101 (.147) | .156 (.051)*** | .091 (.147) |
| N | 3593 | 268 | 3593 | 268 |
| Pseudolog-lik. | -1029.872 | -1012.155 | -1030.045 | -1013.52 |

Note: Redoes Table 5 while restricting the sample between 1945 and 2004 (rather than 1875 and 2004). Heckman probit estimations. The first stage runs a probit estimation of the likelihood that an autocracy breaks down. The second stage runs a probit estimation of the likelihood that an autocracy that has just broken down transitions to democracy as opposed to another authoritarian regime. Models on authoritarian breakdown include the age of the regime, its square and its cube. Robust standard errors clustered by country in parentheses. All explanatory variables are lagged. *** $p < .01$, ** $p < .05$ and * $p < .1$.

Table A19: Effect of Diffusion and Shocks on Authoritarian Breakdowns and Transitions to Democracy (1945-2004; Does Not Control for % Democracies in the World)

| | Spatial Lag 1 | | | | Spatial Lag 2 | | | |
|-----------------------------|-------------------------|----------------------|-------------------------|----------------------|----------------------|-----|-------------------------|----------------------|
| | (1) | (2) | (3) | (4) | (1) | (2) | (3) | (4) |
| % Dem. neigh. | .070 (.110) | .896 (.334)*** | | | 1.311 (.494)*** | | | |
| Ch. % dem. neigh. last year | | | -164 (.306) | 1.895 (.754)** | | | -138 (.511) | .360 (1.296) |
| Hegemonic power volatility | .210 (.065)*** | .237 (.163) | .215 (.066)*** | .237 (.165) | .267 (.162)* | | .215 (.065)*** | .241 (.166) |
| Growth rate | -.016 (.007)** | -.031 (.024) | -.016 (.007)** | -.022 (.024) | -.026 (.025) | | -.016 (.007)** | -.026 (.024) |
| GDP pc (logged) | -.099 (.050)** | .562 (.241)** | -.102 (.048)** | .671 (.240)*** | .560 (.236)** | | -.102 (.048)** | .667 (.246)*** |
| Oil | -5.26e-06 (1.00e-05) | -.0005 (.0002)*** | -5.32e-06 (1.00e-05) | -.0006 (.0002)*** | -.0005 (.0002)*** | | -5.25e-06 (1.00e-05) | -.0006 (.0002)*** |
| Muslim | -.0002 (.002) | -.0006 (.004) | -.0002 (.002) | -.0006 (.004) | -.0004 (.004) | | -.0002 (.002) | -.0003 (.004) |
| Catholic | .004 (.002)** | -.002 (.004) | .004 (.002)** | -.0006 (.004) | -.003 (.004) | | .004 (.002)** | -.0001 (.004) |
| Protestant | -.001 (.003) | -.004 (.007) | -.003 (.003) | -.002 (.008) | -.002 (.007) | | -.003 (.003) | -.002 (.008) |
| Ethnic frac. | -.0002 (.002) | .010 (.004)** | -.0001 (.002) | .009 (.004)** | .010 (.005)** | | -.0001 (.002) | .009 (.004)** |
| Religious frac. | -.002 (.002) | .003 (.006) | -.001 (.002) | .003 (.006) | .003 (.006) | | -.001 (.002) | .003 (.006) |
| Brit. col. | -.104 (.092) | .387 (.285) | -.090 (.091) | .366 (.278) | .453 (.308) | | -.091 (.091) | .349 (.278) |
| # Past trans. | .153 (.053)*** | .154 (.113) | .152 (.053)*** | .143 (.113) | .127 (.119) | | .152 (.053)*** | .145 (.113) |
| N | 3593 | 268 | 3578 | 264 | 268 | | 3578 | 264 |
| Pseudolog-lik. | -1035.965 | -1020.919 | -1035.208 | -1022.484 | -1035.208 | | -1022.484 | -1022.484 |

Note: Redoes Table 5 while restricting the sample between 1945 and 2004 (rather than 1875 and 2004), and without controlling for the % of democracies in the world. Heckman probit estimations. The first stage runs a probit estimation of the likelihood that an autocracy breaks down. The second stage runs a probit estimation of the likelihood that an autocracy that has just broken down transitions to democracy as opposed to another authoritarian regime. Models on authoritarian breakdown include the age of the regime, its square and its cube. Robust standard errors clustered by country in parentheses. All explanatory variables are lagged. *** $p < .01$, ** $p < .05$ and * $p < .1$.

Table A20: Effect of Diffusion and Shocks on Authoritarian Breakdown and Democratization (Excluding Each Region) — Part I

| | EXCLUDED REGION | | | | | |
|----------------------------|-------------------------|----------------------|-------------------------|----------------------|-----------------------|--------------------|
| | Western Countries | | Eastern Europe | | Latin America | |
| | Auth. Break. | Tr. to Dem. | Auth. Break. | Tr. to Dem. | Auth. Break. | Tr. to Dem. |
| % Dem. neigh. | .066 (.115) | .704 (.332)** | .064 (.112) | .626 (.284)** | -.082 (.159) | .566 (.384) |
| % Dem. world | .255 (.268) | 3.427 (.926)*** | .356 (.240) | 1.886 (.797)** | .474 (.313) | 1.839 (.772)** |
| Hegemonic power volatility | .132 (.049)*** | .105 (.135) | .142 (.051)*** | .068 (.129) | .196 (.060)*** | .160 (.157) |
| Growth rate | -.018 (.006)*** | -.020 (.025) | -.019 (.006)*** | -.026 (.020) | -.020 (.007)*** | -.005 (.020) |
| GDP pc (logged) | -.113 (.053)** | .633 (.221)*** | -.100 (.055)* | .563 (.203)*** | -.060 (.057) | .315 (.236) |
| Oil | -3.12e-06 (1.00e-05) | -.0005 (.0002)*** | -3.89e-06 (1.00e-05) | -.0005 (.0002)*** | -1.00e-05 (.00002) | -.002 (.0008)** |
| Muslim | -.0006 (.001) | -.0007 (.004) | .00002 (.001) | -.002 (.003) | -.0005 (.001) | .0002 (.003) |
| Catholic | .003 (.001)** | -.002 (.004) | .004 (.002)** | -.002 (.003) | .0002 (.002) | .0004 (.003) |
| Protestant | -.001 (.003) | -.002 (.007) | -.002 (.002) | .005 (.006) | -.0005 (.002) | .004 (.005) |
| Ethnic frac. | -.0001 (.002) | .008 (.004)* | -.0004 (.002) | .006 (.003)** | .0004 (.002) | .005 (.004) |
| Religious frac. | -.003 (.002)* | .002 (.005) | -.0008 (.002) | -.002 (.005) | -.002 (.002) | -.0002 (.004) |
| Brit. col. | -.092 (.096) | .281 (.335) | -.150 (.088)* | .243 (.263) | -.079 (.100) | .180 (.244) |
| # Past trans. | .169 (.052)*** | .221 (.116)* | .153 (.051)*** | .213 (.082)*** | .156 (.074)** | .205 (.097)** |
| N | 5030 | 394 | 4852 | 405 | 3853 | 210 |
| Pseudolog-lik. | | -1456.376 | | -1484.778 | | -879.846 |

Note: Redoes model 1 of Table 5 of the main text while excluding each region in succession. *Western Countries* are defined as Australia, New Zealand, the United States, Canada as well as countries from Western Europe. Heckman probit estimations. The first stage runs a probit estimation of the likelihood that an autocracy breaks down. The second stage runs a probit estimation of the likelihood that an autocracy that has just broken down transitions to democracy as opposed to another authoritarian regime. Models on authoritarian breakdown include the age of the regime, its square and its cube. Robust standard errors clustered by country in parentheses. All explanatory variables are lagged. *** $p < .01$, ** $p < .05$ and * $p < .1$.

Table A20: Effect of Diffusion and Shocks on Authoritarian Breakdown and Democratization (Excluding Each Region) — Part II

| | EXCLUDED REGION | | | | | |
|----------------------------|-------------------|-------------------|-------------------------|---------------------|-------------------------|---------------------|
| | Middle East | | Sub-Saharan Africa | | Asia | |
| | Auth. Break. | Tr. to Dem. | Auth. Break. | Tr. to Dem. | Auth. Break. | Tr. to Dem. |
| % Dem. neigh. | .033 (.116) | .604 (.260)** | .080 (.125) | .648 (.372)* | .113 (.119) | .383 (.302) |
| % Dem. world | .509 (.265)* | 2.169 (.757)** | .470 (.321) | 2.193 (1.274)* | .213 (.252) | 2.786 (.877)** |
| Hegemonic power volatility | .126 (.051)** | .093 (.119) | .142 (.054)** | -.122 (.162) | .127 (.049)** | .040 (.140) |
| Growth rate | -.018 (.006)** | -.024 (.020) | -.017 (.008)** | -.048 (.026)* | -.017 (.006)** | -.028 (.022) |
| GDP pc (logged) | -.045 (.059) | .600 (.182)** | -.058 (.075) | 1.017 (.197)** | -.108 (.058)* | .701 (.237)** |
| Oil | -.0001 (.0001) | -.0002 (.0002) | -1.54e-06 (1.00e-05) | -.0005 (.0002)** | -5.02e-06 (1.00e-05) | -.0005 (.0002)** |
| Muslim | -.0006 (.001) | .002 (.004) | -.0008 (.002) | -.003 (.004) | -.003 (.002) | -.002 (.006) |
| Catholic | .003 (.001)** | -.002 (.003) | .003 (.002)* | -.007 (.003)* | .0007 (.003) | -.001 (.005) |
| Protestant | -.0006 (.002) | .003 (.005) | .004 (.003)* | .008 (.008) | -.003 (.003) | .006 (.007) |
| Ethnic frac. | .0003 (.002) | .004 (.003) | .001 (.002) | .004 (.006) | .0004 (.002) | .006 (.004) |
| Religious frac. | -.004 (.002)** | -.002 (.004) | -.006 (.002)** | .005 (.006) | -.005 (.002)** | .002 (.006) |
| Brit. col. | -.068 (.108) | .309 (.267) | -.250 (.140)* | .068 (.636) | -.027 (.097) | .277 (.296) |
| # Past trans. | .131 (.050)** | .139 (.075)* | .151 (.061)** | -.071 (.143) | .151 (.051)** | .157 (.101) |
| N | 4775 | 396 | 4121 | 325 | 4844 | 386 |
| Pseudolog-lik. | | -1463.717 | | -1187.116 | | -1430.465 |

Note: Redoes model 1 of Table 5 of the main text while excluding each region in succession. Heckman probit estimations. The first stage runs a probit estimation of the likelihood that an autocracy breaks down. The second stage runs a probit estimation of the likelihood that an autocracy that has just broken down transitions to democracy as opposed to another authoritarian regime. Models on authoritarian breakdown include the age of the regime, its square and its cube. Robust standard errors clustered by country in parentheses. All explanatory variables are lagged. *** $p < .01$, ** $p < .05$ and * $p < .1$.

Table A21: Effect of Diffusion and Shocks on Authoritarian Breakdown and Democratization (Excluding Each Region; Does Not Control for % Democracies in the World) — Part I

| | EXCLUDED REGION | | | | | |
|----------------------------|-------------------------|---------------------|-------------------------|---------------------|----------------------|--------------------|
| | Western Countries | | Eastern Europe | | Latin America | |
| | Auth. Break. | Tr. to Dem. | Auth. Break. | Tr. to Dem. | Auth. Break. | Tr. to Dem. |
| % Dem. neigh. | .101 (.106) | 1.161 (.315)*** | .106 (.107) | .832 (.264)*** | -.034 (.155) | .731 (.373)** |
| Hegemonic power volatility | .141 (.049)*** | .194 (.120) | .157 (.051)*** | .125 (.120) | .207 (.061)*** | .207 (.150) |
| Growth rate | -.019 (.006)*** | -.040 (.024)* | -.020 (.006)*** | -.034 (.019)* | -.021 (.007)*** | -.012 (.019) |
| GDP pc (logged) | -.107 (.052)** | .586 (.219)*** | -.095 (.056)* | .524 (.199)*** | -.053 (.058) | .296 (.216) |
| Oil | -3.35e-06 (1.00e-05) | -.0004 (.0002)** | -3.81e-06 (1.00e-05) | -.0004 (.0002)** | -1.00e-05 (.0002) | -.002 (.0007)** |
| Muslim | -.0006 (.001) | -.001 (.003) | .00002 (.001) | -.002 (.003) | -.0004 (.001) | .00008 (.003) |
| Catholic | .003 (.001)** | -.004 (.003) | .003 (.001)** | -.003 (.003) | -.00004 (.002) | -.0005 (.003) |
| Protestant | -.0007 (.003) | -.0006 (.006) | -.002 (.002) | .004 (.005) | -.0007 (.002) | .002 (.004) |
| Ethnic frac. | .00002 (.001) | .009 (.004)** | -.0002 (.002) | .007 (.003)** | .0009 (.001) | .006 (.003)* |
| Religious frac. | -.003 (.002) | .004 (.005) | .0009 (.002) | -.001 (.004) | -.002 (.002) | .001 (.004) |
| Brit. col. | -.086 (.096) | .330 (.275) | -.136 (.088) | .300 (.240) | -.062 (.100) | .242 (.210) |
| # Past trans. | .173 (.052)*** | .242 (.096)** | .160 (.051)*** | .239 (.076)*** | .164 (.074)** | .230 (.092)** |
| N | 5030 | 394 | 4852 | 405 | 3853 | 210 |
| Pseudolog-lik. | | -1463.163 | | -1488.004 | | -883.349 |

Note: Redoes model 1 of Table 5 of the main text while excluding each region in succession, and without controlling for the % of democracies in the world. *Western Countries* are defined as Australia, New Zealand, the United States, Canada as well as countries from Western Europe. Heckman probit estimations. The first stage runs a probit estimation of the likelihood that an autocracy breaks down. The second stage runs a probit estimation of the likelihood that an autocracy that has just broken down transitions to democracy as opposed to another authoritarian regime. Models on authoritarian breakdown include the age of the regime, its square and its cube. Robust standard errors clustered by country in parentheses. All explanatory variables are lagged. *** $p < .01$, ** $p < .05$ and * $p < .1$.

Table A21: Effect of Diffusion and Shocks on Authoritarian Breakdown and Democratization (Excluding Each Region; Does Not Control for % Democracies in the World) — Part II

| | EXCLUDED REGION | | | | | |
|----------------------------|--------------------|-------------------|-------------------------|---------------------|-------------------------|----------------------|
| | Middle East | | Sub-Saharan Africa | | Asia | |
| | Auth. Break. | Tr. to Dem. | Auth. Break. | Tr. to Dem. | Auth. Break. | Tr. to Dem. |
| % Dem. neigh. | .095 (.111) | .852 (.241)*** | .134 (.122) | .856 (.360)** | .143 (.113) | .794 (.289)*** |
| Hegemonic power volatility | .146 (.051)*** | .153 (.112) | .163 (.054)*** | -.085 (.153) | .135 (.049)*** | .118 (.131) |
| Growth rate | -.020 (.006)*** | -.032 (.019) | -.017 (.009)** | -.050 (.028)* | -.017 (.006)*** | -.038 (.022)* |
| GDP pc (logged) | -.041 (.060) | .566 (.177)*** | -.037 (.071) | 1.010 (.217)*** | -.106 (.058)* | .620 (.229)*** |
| Oil | -.00008 (.0001) | -.0001 (.0002) | -2.95e-06 (1.00e-05) | -.0004 (.0002)** | -4.94e-06 (1.00e-05) | -.0004 (.0002)*** |
| Muslim | -.0004 (.001) | .002 (.003) | -.0009 (.002) | -.003 (.004) | -.003 (.002) | -.003 (.005) |
| Catholic | .002 (.001)* | -.004 (.003) | .002 (.001)* | -.008 (.003)*** | .0004 (.003) | -.004 (.005) |
| Protestant | -.0007 (.002) | .002 (.005) | .004 (.002) | .005 (.007) | -.003 (.003) | .003 (.006) |
| Ethnic frac. | .0006 (.002) | .005 (.003)* | .001 (.002) | .004 (.006) | .0005 (.002) | .006 (.003)* |
| Religious frac. | -.004 (.002)** | -.001 (.004) | -.006 (.002)*** | .006 (.006) | -.006 (.002)** | .003 (.006) |
| Brit. col. | -.047 (.108) | .399 (.235)* | -.235 (.141)* | .253 (.583) | -.019 (.098) | .360 (.256) |
| # Past trans. | .140 (.050)*** | .161 (.069)** | .152 (.060)** | -.088 (.143) | .155 (.051)*** | .192 (.091)** |
| N | 4775 | 396 | 4121 | 325 | 4844 | 386 |
| Pseudolog-lik. | | -1468.73 | | -1189.564 | | -1435.894 |

Note: Redoes model 1 of Table 5 of the main text while excluding each region in succession, and without controlling for the % of democracies in the world. Heckman probit estimations. The first stage runs a probit estimation of the likelihood that an autocracy breaks down. The second stage runs a probit estimation of the likelihood that an autocracy that has just broken down transitions to democracy as opposed to another authoritarian regime. Models on authoritarian breakdown include the age of the regime, its square and its cube. Robust standard errors clustered by country in parentheses. All explanatory variables are lagged. *** $p < .01$, ** $p < .05$ and * $p < .1$.

Table A22: Effect of Diffusion and Shocks on Authoritarian Breakdowns and Transitions to Democracy (Benchmarked Growth Rates)

| | Spatial Lag 1 | | Spatial Lag 2 | | | | | |
|-----------------------------|-------------------------|---------------------------------|----------------------------------|-----------------------------------|----------------------------------|---------------------------------|----------------------------------|----------------------------------|
| | (1) | (2) | (3) | (4) | | | | |
| % Dem. neigh. | .051 (.109) | Tr. to Dem. .603 (.290)** | Auth. Break. -0.254 (.250) | Tr. to Dem. 1.682 (.622)*** | Auth. Break. -0.026 (.160) | Tr. to Dem. .781 (.393)** | Auth. Break. -0.275 (.502) | Tr. to Dem. -0.302 (1.029) |
| Ch. % dem. neigh. last year | | | | | | | | |
| % Dem. world | .327 (.242) | 2.452 (.806)*** | .381 (.238) | 3.230 (.838)*** | .377 (.251) | 2.293 (.841)*** | .383 (.240) | 3.244 (.832)*** |
| Hegemonic power volatility | .147 (.048)*** | .071 (.130) | .155 (.049)*** | .051 (.138) | .146 (.048)*** | .075 (.127) | .157 (.049)*** | .064 (.133) |
| Bench. Growth rate | -.020 (.006)*** | -.022 (.022) | -.020 (.006)*** | -.018 (.023) | -.020 (.006)*** | -.020 (.022) | -.020 (.006)*** | -.021 (.022) |
| GDP pc (logged) | -.085 (.053) | .619 (.203)*** | -.083 (.052) | .824 (.224)*** | -.076 (.056) | .614 (.208)*** | -.083 (.052) | .810 (.227)*** |
| Oil | -6.64e-06 (1.00e-05) | -.0005 (.0002)*** | -7.10e-06 (1.00e-05) | -.0006 (.0002)*** | -8.19e-06 (1.00e-05) | -.0005 (.0002)*** | -7.04e-06 (1.00e-05) | -.0006 (.0002)*** |
| Muslim | -.0007 (.001) | -.001 (.003) | -.0006 (.001)** | -.001 (.004) | -.0007 (.001) | -.001 (.003) | -.0006 (.001) | -.0007 (.004) |
| Catholic | .003 (.001)** | -.002 (.003) | .003 (.001)** | -.002 (.004) | .003 (.001)** | -.002 (.003) | .003 (.001)** | -.002 (.003) |
| Protestant | -.0003 (.002) | .004 (.005) | -.0008 (.002) | .009 (.007) | -.0003 (.002) | .005 (.005) | -.0008 (.002) | .009 (.006) |
| Ethnic frac. | .0003 (.002) | .005 (.003)* | .0004 (.002) | .006 (.004) | .0002 (.002) | .006 (.003)* | .0004 (.002) | .006 (.003)* |
| Religious frac. | -.003 (.002)* | .0005 (.004) | -.003 (.002)* | .0009 (.005) | -.003 (.002)* | .0005 (.004) | -.003 (.002)* | .002 (.005) |
| Brit. col. | -.098 (.092) | .221 (.274) | -.094 (.093) | .242 (.288) | -.100 (.093) | .253 (.282) | -.095 (.093) | .219 (.282) |
| # Past trans. | .150 (.049)*** | .172 (.087)** | .154 (.049)*** | .134 (.105) | .153 (.049)*** | .154 (.087)* | .153 (.049)*** | .142 (.101) |
| N | 5488 | 423 | 5458 | 416 | 5488 | 423 | 5458 | 416 |
| Pseudolog-lik. | | -1584.814 | | -1557.662 | | -1584.697 | | -1559.458 |

Note: Redoes Table 5 with growth rates benchmarked against the world growth rate. Heckman probit estimations. The first stage runs a probit estimation of the likelihood that an autocracy breaks down. The second stage runs a probit estimation of the likelihood that an autocracy that has just broken down transitions to democracy as opposed to another authoritarian regime. Models on authoritarian breakdown include the age of the regime, its square and its cube. Robust standard errors clustered by country in parentheses. All explanatory variables are lagged. *** $p < .01$, ** $p < .05$ and * $p < .1$.

Table A23: Effect of Diffusion and Shocks on Authoritarian Breakdowns and Transitions to Democracy (Excluding Outliers)

| | Spatial Lag 1 | | Spatial Lag 2 | | | | | | |
|-----------------------------|-------------------------|----------------------|----------------------|-------------------------|----------------------|-------------|-------------|-------------------------|----------------------|
| | (1) | (2) | (3) | (4) | | | | | |
| % Dem. neigh. | .059 (.109) | .626 (.295)** | 1.701 (.625)*** | -.002 (.160) | .810 (.396)** | Tr. to Dem. | Tr. to Dem. | Auth. Break. | Tr. to Dem. |
| Ch. % dem. neigh. last year | | | | | | | | | |
| % Dem. world | .386 (.242) | 2.463 (.804)*** | 3.266 (.828)*** | .424 (.250)* | 2.301 (.842)*** | | | -.288 (.506) | -320 (1.041) |
| Hegemonic power volatility | .137 (.048)*** | .059 (.129) | .040 (.136) | .137 (.048)*** | .065 (.126) | | | .445 (.241)* | 3.289 (.824)*** |
| Growth rate | -.018 (.007)*** | -.025 (.021) | -.017 (.022) | -.018 (.007)*** | -.022 (.021) | | | .147 (.049)*** | .052 (.132) |
| Oil | -5.98e-06 (1.00e-05) | -.0005 (.0002)*** | -.0006 (.0002)*** | -6.64e-06 (1.00e-05) | -.0005 (.0002)*** | | | -.018 (.007)*** | -.020 (.021) |
| GDP pc (logged) | -.090 (.054)* | .637 (.205)*** | .836 (.224)*** | -7.29e-06 (1.00e-05) | -.0005 (.0002)*** | | | -6.59e-06 (1.00e-05) | -.0006 (.0002)*** |
| Muslim | -.0007 (.001) | -.001 (.003) | -.001 (.004) | -.083 (.056) | -.629 (.210)*** | | | -.086 (.052)* | .823 (.228)*** |
| Catholic | .003 (.001)** | -.002 (.003) | -.002 (.004) | -.0007 (.001) | -.001 (.003) | | | -.0006 (.001) | -.0007 (.004) |
| Protestant | -.0001 (.002) | .004 (.005) | .009 (.006) | .003 (.001)** | -.002 (.003) | | | .003 (.001)** | -.002 (.004) |
| Ethnic frac. | .0004 (.002) | .006 (.003)* | .006 (.004) | -.0005 (.002) | .005 (.005) | | | -.0005 (.002) | .009 (.006) |
| Religious frac. | -.003 (.002)* | .0004 (.004) | .0009 (.005) | .004 (.002) | .006 (.003)* | | | -.0004 (.002) | .006 (.003)* |
| Brit. col. | -.105 (.091) | .228 (.278) | .246 (.292) | -.101 (.092) | .260 (.285) | | | -.004 (.002)* | .001 (.005) |
| # Past trans. | .151 (.050)*** | .169 (.089)* | .132 (.107) | .153 (.049)*** | .151 (.089)* | | | -.102 (.092) | .225 (.286) |
| N | 5453 | 423 | 416 | 5423 | 423 | | | .154 (.050)*** | .140 (.104) |
| Pseudolog-lik. | | -1583.243 | -1556.217 | | -1583.202 | | | 5423 | 416 |

Note: Redoes Table 5 without observations with growth rates below -20 percent or above 20 percent. Heckman probit estimations. The first stage runs a probit estimation of the likelihood that an autocracy breaks down. The second stage runs a probit estimation of the likelihood that an autocracy that has just broken down transitions to democracy as opposed to another authoritarian regime. Models on authoritarian breakdown include the age of the regime, its square and its cube. Robust standard errors clustered by country in parentheses. All explanatory variables are lagged. *** $p < .01$, ** $p < .05$ and * $p < .1$.

Table A25: Effect of Diffusion on Democratization (Country Fixed Effects)

| | Spatial Lag 1 | | | Spatial Lag 2 | | | | |
|--------------------------------|---------------------|---------------------|---------------------|---------------------|---------------------|---------------------|---------------------|---------------------|
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| % Dem. neigh. | .510 (.337) | | | | 1.263 (.587)** | | | |
| Ch. % dem. neigh. last year | | .405 (.405) | | | | .241 (.809) | | |
| Ch. % dem. neigh. last 2 years | | | .106 (.295) | | | | .137 (.691) | |
| Ch. % dem. neigh. last 3 years | | | | -.155 (.278) | | | | -.114 (.564) |
| % Dem. world | 5.582 (1.246)*** | 5.910 (1.161)*** | 5.868 (1.173)*** | 5.920 (1.169)*** | 5.000 (1.341)*** | 5.921 (1.164)*** | 5.868 (1.179)*** | 5.908 (1.176)*** |
| Hegemonic power volatility | .056 (.123) | .041 (.126) | .055 (.125) | .055 (.125) | .057 (.123) | .041 (.126) | .053 (.128) | .055 (.126) |
| Growth rate | -.073 (.018)*** | -.077 (.018)*** | -.078 (.018)*** | -.080 (.018)*** | -.071 (.018)*** | -.077 (.018)*** | -.078 (.018)*** | -.080 (.018)*** |
| GDP pc (logged) | 1.562 (.318)*** | 1.604 (.304)*** | 1.628 (.311)*** | 1.614 (.310)*** | 1.544 (.322)*** | 1.604 (.306)*** | 1.626 (.313)*** | 1.614 (.310)*** |
| Oil | .002 (.0008)* | .001 (.0008)* | .002 (.0008)** | .002 (.0008)** | .002 (.0008)* | .002 (.0008)* | .002 (.0008)** | .002 (.0008)** |
| # Past trans. | -.657 (.192)*** | -.623 (.198)*** | -.629 (.200)*** | -.621 (.204)*** | -.736 (.177)*** | -.619 (.199)*** | -.627 (.200)*** | -.624 (.204)*** |
| N | 3386 | 3359 | 3322 | 3289 | 3386 | 3359 | 3322 | 3289 |
| Pseudolog-lik. | -359.033 | -357.82 | -355.115 | -356.475 | -356.173 | -358.068 | -355.13 | -356.551 |

Note: Redoes Table 4 with country fixed effects. Dynamic probit estimations. Robust standard errors clustered by country in parentheses. All explanatory variables are lagged. *** $p < .01$, ** $p < .05$ and * $p < .1$.

Table A26: Effect of Alternative Political Shocks on Authoritarian Breakdowns and Transitions to Democracy

(2)

| | Auth. Break. | Tr. to Dem. | Auth. Break. | Tr. to Dem. |
|-----------------------------|-------------------------|----------------------|-------------------------|----------------------|
| % Dem. neigh. | -.014 (.119) | .572 (.304)* | -.122 (.121) | .503 (.316) |
| % Dem. world | 1.212 (.525)** | 3.214 (1.245)*** | .471 (.258)* | 2.168 (.904)** |
| Hegemonic Power Volatility | .134 (.050)*** | .111 (.129) | .125 (.050)** | .072 (.127) |
| Growth rate | -.016 (.006)** | -.027 (.021) | -.018 (.006)*** | -.029 (.021) |
| GDP pc (logged) | -.081 (.052) | .633 (.209)*** | -.105 (.048)** | .578 (.210)*** |
| Oil | -7.53e-06 (1.00e-05) | -.0005 (.0002)*** | -6.26e-06 (1.00e-05) | -.0005 (.0002)*** |
| Muslim | -.0006 (.001) | -.001 (.003) | -.0002 (.001) | -.0008 (.003) |
| Catholic | .003 (.001)** | -.003 (.003) | .002 (.001)* | -.003 (.003) |
| Protestant | -.0002 (.002) | .004 (.006) | -.001 (.002) | .003 (.006) |
| Ethnic frac. | .0004 (.001) | .006 (.004) | .0002 (.001) | .005 (.003) |
| Religious frac. | -.004 (.002)** | .0002 (.005) | -.003 (.002) | .002 (.004) |
| Brit. col. | -.127 (.093) | .126 (.323) | -.103 (.093) | .133 (.317) |
| # Past trans. | .153 (.055)*** | .187 (.093)** | .124 (.052)** | .148 (.092) |
| Polarized Order | -.074 (.124) | -.182 (.258) | | |
| Democratic Order | -.251 (.196) | -.409 (.355) | | |
| Alliance with France | | | | |
| Alliance with Germany | | | .110 (.403) | 5.662 (.886)*** |
| Alliance with Japan | | | -.538 (.291)* | .889 (.824) |
| Alliance with Russia/USSR | | | .711 (.278)** | -.241 (.930) |
| Alliance with US | | | -.219 (.101)** | -.542 (.654) |
| Alliance with US (Cold War) | | | .265 (.253) | .112 (.396) |
| Alliance with UK | | | .049 (.247) | .246 (.394) |
| | | | -.192 (.261) | -4.500 (.952)*** |
| N | 5229 | 413 | 5229 | 413 |
| Pseudolog-lik. | | -1541.107 | | -1529.542 |

Note: Redoes model 1 of Table 5 with additional measures of political shocks. Heckman probit estimations. The first stage runs a probit estimation of the likelihood that an autocracy breaks down. The second stage runs a probit estimation of the likelihood that an autocracy that has just broken down transitions to democracy as opposed to another authoritarian regime. *Polarized Order* takes the value 1 in 1933-90, and *Democratic Order* takes the value 1 in 1919-32 and 1991-2000. These variables and the variables on alliances are taken from Boix (2011). Models on authoritarian breakdown include the age of the regime, its square and its cube. Robust standard errors clustered by country in parentheses. All explanatory variables are lagged. *** $p < .01$, ** $p < .05$ and * $p < .1$.

Table A27: Effect of Diffusion and Shocks on Authoritarian Breakdowns and Transitions to Democracy (US Power Volatility)

| | Spatial Lag 1 | | | Spatial Lag 2 | | |
|-----------------------------|-------------------------|---------------------|-------------------------|---------------------|-------------------------|---------------------|
| | (1) | (2) | (3) | (4) | (5) | (6) |
| | Auth. Break. | Tr. to Dem. | Auth. Break. | Tr. to Dem. | Auth. Break. | Tr. to Dem. |
| % Dem. neigh. | .026 (.117) | .663 (.318)** | | .820 (.420)* | | |
| Ch. % dem. neigh. last year | | | -288 (.255) | 1,790 (.608)** | -457 (.517) | -379 (1.087) |
| % Dem. world | .411 (.238)* | 1,796 (.851)** | 435 (.231)* | 2,618 (.868)** | 439 (.232)* | 2,670 (.850)** |
| US Power Volatility | 2,202 (.740)** | 3,523 (2.050)* | 2,394 (.756)** | 2,651 (2.108) | 2,460 (.749)** | 2,831 (2.061) |
| Growth rate | -.016 (.006)** | -.029 (.022) | -.016 (.006)** | -.020 (.023) | -.017 (.006)** | -.023 (.022) |
| GDP pc (logged) | -.082 (.053) | .668 (.198)** | -.082 (.052) | .856 (.209)** | -.081 (.052) | .849 (.217)** |
| Oil | -7.12e-06 (1.00e-05) | -.0005 (.0002)** | -7.01e-06 (1.00e-05) | -.0006 (.0002)** | -8.64e-06 (1.00e-05) | -.0006 (.0002)** |
| Muslim | -.0007 (.001) | -.0009 (.003) | -.0006 (.001) | -.0007 (.004) | -.0005 (.001) | -.0005 (.004) |
| Catholic | .003 (.001)** | -.003 (.003) | .003 (.001)** | -.003 (.004) | .003 (.001)** | -.003 (.004) |
| Protestant | -.0004 (.002) | .005 (.006) | -.001 (.002) | .009 (.007) | -.0009 (.002) | .009 (.007) |
| Ethnic frac. | .0003 (.001) | .006 (.003) | .0004 (.002) | .006 (.004) | .0004 (.002) | .006 (.004)* |
| Religious frac. | -.004 (.002)** | -.00007 (.005) | -.004 (.002)** | .0005 (.005) | -.004 (.002)** | -.0003 (.005) |
| Brit. col. | -.129 (.092) | .135 (.329) | -.126 (.093) | .152 (.342) | -.127 (.093) | .133 (.335) |
| # Past trans. | .150 (.050)** | .188 (.099)* | .153 (.050)** | .147 (.114) | .152 (.049)** | .153 (.112) |
| N | 5488 | 423 | 5458 | 416 | 5488 | 416 |
| Pseudolog-lik. | -1547.687 | -1521.646 | -1521.646 | -1547.632 | -1523.317 | -1523.317 |

Note: Redoes Table 5 but measures political shocks with the absolute value of the change in the US share of hegemonic power over the last three years, rather than the average of the absolute value of the change in the share of power of each hegemon over the last three years. Heckman probit estimations. The first stage runs a probit estimation of the likelihood that an autocracy breaks down. The second stage runs a probit estimation of the likelihood that an autocracy that has just broken down transitions to democracy as opposed to another authoritarian regime. Models on authoritarian breakdown include the age of the regime, its square and its cube. Robust standard errors clustered by country in parentheses. All explanatory variables are lagged. *** $p < .01$, ** $p < .05$ and * $p < .1$.

Table A28: Effect of Diffusion on Democratization (Interaction Between Diffusion and Economic Shocks)

| | Spatial Lag 1 | | | Spatial Lag 2 | | | | |
|--|--------------------|--------------------|--------------------|--------------------|-------------------|--------------------|--------------------|--------------------|
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| % Dem. neigh. | .381 (.152)** | | | | .465 (.229)** | | | |
| Ch. % dem. neigh. last year | | .559 (.346) | | | | .667 (.601) | | |
| Ch. % dem. neigh. last 2 years | | | .290 (.252) | | | | .402 (.461) | |
| Ch. % dem. neigh. last 3 years | | | | .083 (.220) | | | | .364 (.426) |
| % Dem. world | 1.111 (.422)*** | 1.422 (.401)*** | 1.442 (.401)*** | 1.429 (.406)*** | 1.061 (.430)** | 1.411 (.403)*** | 1.445 (.399)*** | 1.404 (.405)*** |
| Hegemonic power volatility | .184 (.076)** | .155 (.078)** | .154 (.079)* | .167 (.079)** | .183 (.075)** | .154 (.077)** | .147 (.079)* | .157 (.079)** |
| Growth rate | -.034 (.014)** | -.023 (.009)** | -.021 (.009)** | -.023 (.009)** | -.034 (.016)** | -.023 (.009)** | -.025 (.010)*** | -.027 (.010)*** |
| % Dem. neigh.*Growth rate | .027 (.027) | | | | .024 (.036) | | | |
| Ch. % dem. neigh. last year*Growth rate | | -.080 (.094) | | | | -.077 (.157) | | |
| Ch. % dem. neigh. last 2 years*Growth rate | | | -.049 (.071) | | | | .130 (.099) | |
| Ch. % dem. neigh. last 3 years*Growth rate | | | | -.043 (.060) | | | | .061 (.093) |
| GDP pc (logged) | .231 (.072)*** | .271 (.068)*** | .274 (.068)*** | .275 (.068)*** | .218 (.078)*** | .274 (.068)*** | .275 (.069)*** | .275 (.069)*** |
| Oil | -.0004 (.0002)* | -.0004 (.0002)* | -.0004 (.0002)* | -.0004 (.0002)* | -.0004 (.0002) | -.0004 (.0002)* | -.0004 (.0002)* | -.0004 (.0002)* |
| Muslim | -.001 (.002) | -.001 (.003) | -.001 (.003) | -.001 (.003) | -.001 (.002) | -.001 (.003) | -.001 (.003) | -.001 (.003) |
| Catholic | .0005 (.002) | .001 (.002) | .001 (.002) | .001 (.002) | .0004 (.002) | .001 (.002) | .001 (.002) | .001 (.002) |
| Protestant | .004 (.003) | .004 (.003) | .004 (.003) | .004 (.003) | .004 (.003)* | .004 (.003) | .004 (.003) | .004 (.003)* |
| Ethnic frac. | .002 (.002) | .002 (.002) | .002 (.002) | .002 (.002) | .002 (.002) | .002 (.002) | .002 (.002) | .002 (.002) |
| Religious frac. | -.003 (.003) | -.003 (.003) | -.003 (.003) | -.003 (.003) | -.003 (.003) | -.003 (.003) | -.003 (.003) | -.003 (.003) |
| Brit. col. | .015 (.136) | .004 (.138) | .007 (.139) | .009 (.138) | .050 (.137) | .005 (.139) | .002 (.139) | .005 (.139) |
| # Past trans. | .255 (.046)*** | .262 (.045)*** | .250 (.044)*** | .261 (.044)*** | .249 (.045)*** | .266 (.044)*** | .255 (.043)*** | .263 (.044)*** |
| N | 5621 | 5587 | 5526 | 5459 | 5621 | 5587 | 5526 | 5459 |
| Pseudolik. | -472.111 | -471.016 | -468.238 | -469.806 | -473.048 | -471.586 | -467.942 | -469.535 |

Note: Redoes Table 4 with interaction terms between growth and diffusion. Dynamic probit estimations. Robust standard errors clustered by country in parentheses. All explanatory variables are lagged. *** $p < .01$, ** $p < .05$ and * $p < .1$.

Table A29: Effect of Diffusion and Shocks on Authoritarian Breakdowns and Transitions to Democracy (Interaction Between Diffusion and Economic Shocks)

| | Spatial Lag 1 | | | | Spatial Lag 2 | | | |
|---|-------------------------|----------------------|-------------------------|----------------------|-------------------------|----------------------|-------------------------|----------------------|
| | (1) | (2) | (3) | (4) | (1) | (2) | (3) | (4) |
| | Auth. Break. | Tr. to Dem. | Auth. Break. | Tr. to Dem. | Auth. Break. | Tr. to Dem. | Auth. Break. | Tr. to Dem. |
| % Dem. neigh. | .0001 (.115) | .633 (.303)** | | | -.050 (.166) | .770 (.398)* | | |
| Ch. % dem. neigh. last year | | | -.228 (.252) | 1.789 (.647)*** | | | -.412 (.508) | -.310 (1.085) |
| % Dem. world | .303 (.239) | 2.440 (.821)*** | .403 (.235)* | 3.238 (.832)*** | .355 (.252) | 2.257 (.845)*** | .414 (.238)* | 3.270 (.826)*** |
| Hegemonic power volatility | .142 (.047)*** | .060 (.130) | .148 (.049)*** | .042 (.136) | .143 (.048)*** | .068 (.126) | .149 (.049)*** | .054 (.131) |
| Growth rate | -.034 (.008)*** | -.021 (.033) | -.018 (.006)*** | -.018 (.022) | -.033 (.010)*** | -.036 (.035) | -.019 (.006)*** | -.021 (.021) |
| % Dem. neigh.*Growth rate | .063 (.025)** | -.008 (.076) | | | .051 (.031) | .045 (.096) | | |
| Ch. % dem. neigh. last year*Growth rate | | | -.031 (.081) | -.077 (.154) | | | .095 (.129) | -.019 (.334) |
| GDP pc (logged) | -.087 (.053)* | .634 (.207)*** | -.082 (.052) | .831 (.224)*** | -.079 (.055) | .627 (.208)*** | -.081 (.052) | .822 (.226)*** |
| Oil | -4.85e-06 (1.00e-05) | -.0005 (.0002)*** | -7.33e-06 (1.00e-05) | -.0006 (.0002)*** | -7.16e-06 (1.00e-05) | -.0005 (.0002)*** | -7.38e-06 (1.00e-05) | -.0006 (.0002)*** |
| Muslim | -.0008 (.001) | -.001 (.003) | -.0006 (.001) | -.001 (.004) | -.0007 (.001) | -.001 (.003) | -.0006 (.001) | -.0008 (.004) |
| Catholic | .003 (.001)** | -.002 (.003) | .003 (.001)** | -.002 (.004) | .003 (.001)** | -.002 (.003) | .003 (.001)** | -.002 (.004) |
| Protestant | -.0006 (.002) | .004 (.006) | -.0008 (.002) | .009 (.007) | -.0004 (.002) | .005 (.005) | -.0008 (.002) | .009 (.006) |
| Ethnic frac. | .0003 (.001) | .006 (.003)* | .0005 (.001) | .006 (.004) | .0003 (.002) | .006 (.003)* | .0005 (.002) | .006 (.003)* |
| Religious frac. | -.003 (.002)* | .0005 (.004) | -.003 (.002)* | .0008 (.005) | -.003 (.002)* | .0005 (.004) | -.003 (.002)* | .00006 (.005) |
| Brit. col. | -.096 (.091) | .246 (.282) | -.093 (.093) | .249 (.291) | -.096 (.093) | .258 (.287) | -.095 (.093) | .229 (.286) |
| # Past trans. | .156 (.048)*** | .171 (.090)* | .156 (.050)*** | .134 (.107) | .155 (.049)*** | .152 (.087)* | .156 (.049)*** | .141 (.103) |
| N | 5488 | 423 | 5458 | 416 | 5488 | 423 | 5458 | 416 |
| Pseudolog-lik. | | -1581.479 | | -1558 | | -1583.617 | | -1559.65 |

Note: Redoes Table 5 with interaction terms between growth and diffusion. Heckman probit estimations. The first stage runs a probit estimation of the likelihood that an autocracy breaks down. The second stage runs a probit estimation of the likelihood that an autocracy that has just broken down transitions to democracy as opposed to another authoritarian regime. Models on authoritarian breakdown include the age of the regime, its square and its cube. Robust standard errors clustered by country in parentheses. All explanatory variables are lagged. *** $p < .01$, ** $p < .05$ and * $p < .1$.

Table A30: Effect of Diffusion and Shocks on Authoritarian Breakdowns and Transitions to Democracy (Change in the Number of Democratic Neighbors)

| | Spatial Lag 1 (1) | | Spatial Lag 2 (2) | |
|-----------------------------|-------------------------|---------------------|-------------------------|---------------------|
| | Auth. Break. | Tr. to Dem. | Auth. Break. | Tr. to Dem. |
| Ch. # dem. neigh. last year | .054 (.065) | .304 (.095)** | .018 (.029) | -.084 (.080) |
| % Dem. world | .409 (.241)* | 3.091 (.817)** | .408 (.243)* | 3.197 (.817)** |
| Hegemonic power volatility | .153 (.048)** | .064 (.130) | .152 (.048)** | .093 (.127) |
| Growth rate | -.019 (.006)** | -.023 (.021) | -.019 (.006)** | -.025 (.020) |
| GDP pc (logged) | -.075 (.050) | .809 (.220)** | -.075 (.050) | .822 (.224)** |
| Oil | -7.96e-06 (1.00e-05) | -.0006 (.0002)** | -8.04e-06 (1.00e-05) | -.0006 (.0002)** |
| Muslim | -.0008 (.001) | -.002 (.004) | -.0008 (.001) | -.002 (.003) |
| Catholic | .003 (.001)* | -.003 (.003) | .003 (.001)* | -.003 (.003) |
| Protestant | -.0009 (.002) | .007 (.006) | -.0009 (.002) | .008 (.006) |
| Ethnic frac. | .0002 (.001) | .005 (.003) | .0002 (.001) | .005 (.003) |
| Religious frac. | -.003 (.002)* | -.0002 (.005) | -.003 (.002)* | -.0008 (.005) |
| Brit. col. | -.103 (.093) | .248 (.285) | -.103 (.093) | .213 (.281) |
| # Past trans. | .156 (.049)** | .155 (.100) | .157 (.049)** | .158 (.101) |
| N | 5412 | 5412 | 5412 | 5412 |
| Pseudolog-lik. | | -1553.449 | | -1554.508 |

Note: Redoes models 2 and 4 of Table 5 with a variable indicating the change over the last year in the number (rather than proportion) of neighbors that are democratic. Heckman probit estimations. The first stage runs a probit estimation of the likelihood that an autocracy breaks down. The second stage runs a probit estimation of the likelihood that an autocracy that has just broken down transitions to democracy as opposed to another authoritarian regime. Models on authoritarian breakdown include the age of the regime, its square and its cube. Robust standard errors clustered by country in parentheses. *** $p < .01$, ** $p < .05$ and * $p < .1$.

Table A31: Instrumental Variable Estimations of the Effect of Growth on Authoritarian Breakdowns and Transitions to Democracy

| | (1) | | (2) | |
|------------------------------|------------------------|----------------------|----------------------|----------------------|
| | Auth. Break. | Tr. to Dem. | Auth. Break. | Tr. to Dem. |
| % Dem. neigh. | -.021 (.125) | .509 (.389) | | |
| Ch. % dem. neigh. last year | | | -.143 (.341) | 1.516 (.862)* |
| % Dem. world | -2.497 (.871)*** | 5.980 (1.986)*** | -2.394 (.882)*** | 6.265 (1.754)*** |
| Hegemonic power volatility | .020 (.129) | .631 (.289)** | .043 (.129) | .525 (.270)* |
| Growth Rate | -.413 (.129)*** | .449 (.401) | -.396 (.130)*** | .448 (.355) |
| GDP pc (logged) | .330 (.155)** | .167 (.594) | .296 (.155)* | .271 (.583) |
| Oil | 1.00e-05 (1.00e-05) | -.0005 (.0002)* | .00002 (1.00e-05) | -.0005 (.0003)** |
| Muslim | -.004 (.002)* | .004 (.006) | -.004 (.002)* | .004 (.005) |
| Catholic | -.0006 (.002) | .002 (.005) | -.0002 (.002) | .001 (.004) |
| Protestant | -.011 (.004)*** | .006 (.013) | -.012 (.004)*** | .011 (.012) |
| Ethnic frac. | -.008 (.003)*** | .017 (.007)** | -.007 (.003)*** | .016 (.006)** |
| Religious frac. | .002 (.002) | -.004 (.005) | .002 (.002) | -.003 (.005) |
| Brit. col. | .086 (.125) | .248 (.403) | .091 (.123) | .223 (.394) |
| # Past trans. | .248 (.060)*** | -.096 (.218) | .241 (.060)*** | -.125 (.188) |
| Growth as Dependent Variable | | | | |
| Rain dev. | | .177 (.054)*** | | .179 (.054)*** |
| % Dem. neigh. | | -.122 (.317) | | |
| Ch. % dem. neigh. last year | | | | -.488 (.909) |
| % Dem. world. | | -5.711 (.753)*** | | -5.742 (.742)*** |
| GDP pc (logged) | | 1.080 (.110)*** | | 1.072 (.108)*** |
| Oil | | .00005 (.00002)** | | .00005 (.00002)** |
| Muslim | | -.011 (.003)*** | | -.011 (.003)*** |
| Catholic | | -.013 (.003)*** | | -.013 (.003)*** |
| Protestant | | -.021 (.006)*** | | -.021 (.006)*** |
| Ethnic frac. | | -.018 (.003)*** | | -.018 (.003)*** |
| Religious frac. | | .003 (.004) | | .003 (.004) |
| Brit. col. | | .536 (.188)*** | | .541 (.189)*** |
| # Past trans. | | .215 (.125)* | | .211 (.125)* |
| F-stat. instr. | | 10.79*** | | 10.97*** |
| N | 3244 | 241 | 3221 | 238 |
| Pseudolog-lik. | -797.729 | -920.749 | -789.776 | -910.324 |

Note: The first stage regressions use rainfall to instrument for growth. The second stage regressions use Heckman probit models. We run a probit estimation of the likelihood that an autocracy breaks down. We then run a probit estimation of the likelihood that an autocracy that has just broken down transitions to democracy as opposed to another authoritarian regime. Models on authoritarian breakdown include the age of the regime, its square and its cube. Robust standard errors clustered by country in parentheses. All explanatory variables are lagged. *** $p < .01$, ** $p < .05$ and * $p < .1$.

2 Interaction Between Economic and Political Shocks

Conditional relationships also warrant consideration, particularly the possibility that economic shocks become more destabilizing when they occur concurrently with political shocks, and vice versa. In Tables A32 and A33 we redo Tables 4 and 5 with an interaction terms between *Growth rate* and *Hegemonic power volatility*. The interaction term is never statistically significant.

However, one must be cautious when interpreting interaction terms in non-linear models (Ai and Norton, 2003). We thus provide marginal effect plots based on model 1 of Table A33 in Figure A1. All variables are kept at their means. The key variable of interest – growth in the two left panels and *Hegemonic power volatility* in the two right panels – is also kept at its mean.¹

The bottom two panels of Figure A1 suggest that the effect of *Growth rate* and *Hegemonic power volatility* on the choice to democratize after an authoritarian breakdown is never statistically significant, regardless of the value of the conditioning variable. The top two panels, however, do provide some evidence suggesting that the effect of *Growth rate* (*Hegemonic power volatility*) on authoritarian breakdowns is conditional on *Hegemonic power volatility* (*Growth rate*). Specifically, the marginal effect of economic contractions on the probability of autocratic breakdown increases slightly at higher levels of *Hegemonic power volatility*. When there is no secure and dominant world power to bail out client states and preserve order, domestic economic performance matters a little bit more for the survival of autocracies. Similarly, political shocks are more likely to destabilize autocracies that are experiencing economic difficulties.

However, the evidence in favor of an interactive effect is weak. In particular, while the effect of economic contractions strengthens as *Hegemonic power volatility* increases, and

¹We use its mean value for an illustration. The marginal effect of, say, growth may vary at different growth levels in non-linear model interactions.

Table A32: Effect of Diffusion on Democratization (Interaction Between Economic and Political Shocks)

| | Spatial Lag 1 | | | | Spatial Lag 2 | | | |
|--|--------------------|--------------------|--------------------|--------------------|-------------------|--------------------|--------------------|--------------------|
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| % Dem. neigh. | .396 (.152)*** | | | | .472 (.228)** | | | |
| Ch. % dem. neigh. last year | | .534 (.333) | | | | .593 (.569) | | |
| Ch. % dem. neigh. last 2 years | | | .277 (.245) | | | | .509 (.435) | |
| Ch. % dem. neigh. last 3 years | | | | | | | | .391 (.411) |
| % Dem. world | 1.136 (.418)*** | 1.424 (.400)*** | 1.449 (.401)*** | .072 (.212) | 1.083 (.428)** | 1.419 (.401)*** | 1.422 (.396)*** | 1.393 (.401)*** |
| Hegemonic power volatility | .188 (.077)** | .157 (.078)** | .151 (.080)* | .167 (.079)** | .186 (.076)** | .155 (.078)** | .145 (.079)* | .159 (.079)* |
| Growth rate | -.019 (.014) | -.024 (.014)* | -.025 (.014)* | -.024 (.014)* | -.019 (.014) | -.023 (.014)* | -.024 (.014)* | -.024 (.014)* |
| Growth rate*Hegemonic power volatility | -.010 (.016) | -.0009 (.015) | .005 (.014) | -.0008 (.015) | -.009 (.017) | -.001 (.016) | .004 (.015) | -.0008 (.015) |
| GDP pc (logged) | .226 (.072)*** | .273 (.068)*** | .275 (.069)*** | .276 (.068)*** | .218 (.078)*** | .274 (.068)*** | .276 (.069)*** | .276 (.069)*** |
| Oil | -.0004 (.0002) | -.0004 (.0002)* | -.0004 (.0002)* | -.0002 (.0002)* | -.0004 (.0002) | -.0004 (.0002)* | -.0004 (.0002)* | -.0004 (.0002)* |
| Muslim | -.001 (.002) | -.001 (.003) | -.001 (.003) | -.001 (.003) | -.001 (.002) | -.001 (.003) | -.001 (.003) | -.001 (.003) |
| Catholic | .0006 (.002) | .001 (.002) | .001 (.002) | .001 (.002) | .0005 (.002) | .001 (.002) | .001 (.002) | .001 (.002) |
| Protestant | .004 (.003) | .004 (.003) | .004 (.003) | .004 (.003) | .005 (.003)* | .004 (.003) | .004 (.003) | .004 (.003) |
| Ethnic frac. | .002 (.002) | .002 (.002) | .002 (.002) | .002 (.002) | .002 (.002) | .002 (.002) | .002 (.002) | .002 (.002) |
| Religious frac. | -.003 (.003) | -.003 (.003) | -.003 (.003) | -.003 (.003) | -.003 (.003) | -.003 (.003) | -.003 (.003) | -.003 (.003) |
| Brit. col. | .012 (.136) | .002 (.138) | .003 (.139) | .007 (.139) | .049 (.137) | .005 (.137) | .005 (.139) | .007 (.139) |
| # Past trans. | .252 (.047)*** | .264 (.044)*** | .253 (.043)*** | .263 (.044)*** | .246 (.046)*** | .266 (.044)*** | .255 (.043)*** | .262 (.044)*** |
| N | 5621 | 5587 | 5526 | 5459 | 5621 | 5587 | 5526 | 5459 |
| Pseudolog-lik. | -472.325 | -471.239 | -468.388 | -470.003 | -473.089 | -471.668 | -468.338 | -469.691 |

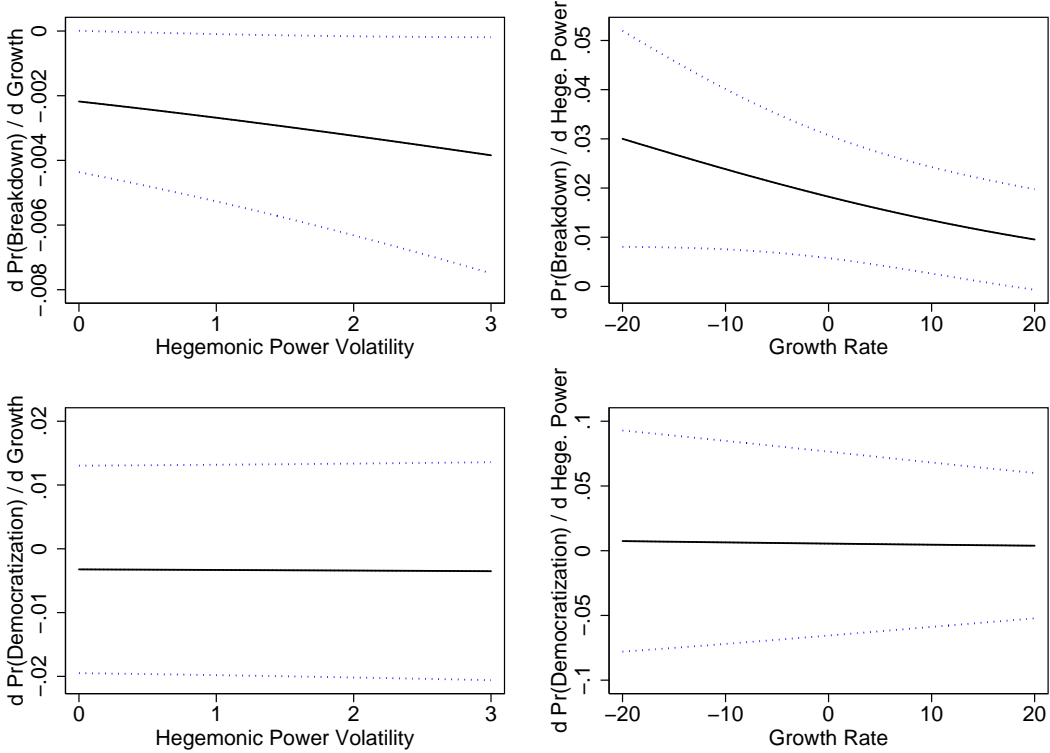
Note: Redoes Table 4 with an interaction term between economic and political shocks. Dynamic probit estimations. Robust standard errors clustered by country in parentheses. All explanatory variables are lagged. *** $p < .01$, ** $p < .05$ and * $p < .1$.

Table A33: Effect of Diffusion and Shocks on Authoritarian Breakdowns and Transitions to Democracy (Interaction Between Economic and Political Shocks)

| | Spatial Lag 1 | | | | Spatial Lag 2 | | | |
|--|-------------------------|---------------------|-------------------------|---------------------|-------------------------|---------------------|-------------------------|---------------------|
| | (1) | (2) | (3) | (4) | (1) | (2) | (3) | (4) |
| | Auth. Break. | Tr. to Dem. | Auth. Break. | Tr. to Dem. | Auth. Break. | Tr. to Dem. | Auth. Break. | Tr. to Dem. |
| % Dem. neigh. | .059 (.109) | .632 (.298)** | -263 (.252) | 1.706 (.620)** | -.009 (.160) | .811 (.398)** | -.300 (.506) | -.308 (1.023) |
| Ch. % dem. neigh. last year | | | | | | | | |
| % Dem. world | .346 (.239) | 2.451 (.813)** | .405 (.234)* | 3.255 (.841)** | .389 (.249) | 2.289 (.848)** | .407 (.236)* | 3.276 (.837)** |
| Hegemonic power volatility | .137 (.049)** | .074 (.129) | .145 (.051)** | .051 (.137) | .136 (.049)** | .078 (.128) | .147 (.051)** | .060 (.133) |
| Growth rate | -.019 (.010)** | -.019 (.028) | -.020 (.010)** | -.013 (.029) | -.019 (.010)** | -.017 (.028) | -.020 (.010)** | -.017 (.028) |
| Growth rate*Hegemonic power volatility | .003 (.014) | -.011 (.031) | .003 (.014) | -.008 (.033) | .003 (.014) | -.010 (.032) | .003 (.014) | -.006 (.031) |
| GDP pc (logged) | -.085 (.053) | .639 (.205)** | -.082 (.052) | .837 (.223)** | -.078 (.055) | .631 (.209)** | -.082 (.052) | .824 (.227)** |
| Oil | -6.68e-06 (1.00e-05) | -.0005 (.0002)** | -7.34e-06 (1.00e-05) | -.0006 (.0002)** | -8.12e-06 (1.00e-05) | -.0005 (.0002)** | -7.29e-06 (1.00e-05) | -.0006 (.0002)** |
| Muslim | -.0017 (.001) | -.001 (.003) | -.0006 (.001) | -.001 (.004) | -.0007 (.001) | -.001 (.003) | -.0006 (.001) | -.0008 (.004) |
| Catholic | .003 (.001)** | -.002 (.003) | .003 (.001)** | -.002 (.004) | .003 (.001)** | -.002 (.003) | .003 (.001)** | -.002 (.004) |
| Protestant | -.0003 (.002) | .004 (.006) | -.0008 (.002) | .009 (.007) | -.0003 (.002) | .005 (.005) | -.0008 (.002) | .009 (.007) |
| Ethnic frac. | .0004 (.002) | .006 (.003)* | .0005 (.002) | .006 (.004) | .0003 (.002) | .006 (.003)* | .0005 (.002) | .006 (.003)* |
| Religious frac. | -.003 (.002)* | .0004 (.004) | -.003 (.002)* | .0008 (.005) | -.003 (.002)* | .004 (.004) | -.003 (.002)* | .001 (.005) |
| Brit. col. | -.097 (.092) | .235 (.279) | -.093 (.093) | .251 (.293) | -.099 (.093) | .267 (.286) | -.094 (.093) | .230 (.287) |
| # Past trans. | .152 (.049)** | .167 (.090)* | .156 (.049)** | .131 (.108) | .155 (.049)** | .150 (.090)* | .155 (.049)** | .139 (.104) |
| N | 5488 | 5488 | 5458 | 5458 | 5488 | 5488 | 5458 | 5458 |
| Pseudolog-lik. | -1585.11 | -1558.057 | -1558.057 | -1585.079 | -1585.079 | -1559.867 | -1559.867 | -1559.867 |

Note: Redoes Table 5 with an interaction term between economic and political shocks. Heckman probit estimations. The first stage runs a probit estimation of the likelihood that an autocracy breaks down. The second stage runs a probit estimation of the likelihood that an autocracy that has just broken down transitions to democracy as opposed to another authoritarian regime. Models on authoritarian breakdown include the age of the regime, its square and its cube. Robust standard errors clustered by country in parentheses. All explanatory variables are lagged. *** $p < .01$, ** $p < .05$ and * $p < .1$.

Figure A1: Marginal Effects of Growth Rate and Hegemonic Power Volatility



Note: Dashed lines give 95 percent confidence intervals. Based on model 1 of Table A33.

vice versa, the confidence intervals also become much wider and they overlap highly, which suggests that these changes are not statistically significant. In fact, whether *Growth rate* (or *Hegemonic power volatility*) is statistically significant or not does not depend on *Hegemonic power volatility* (*Growth rate*). The effect of economic contractions on authoritarian breakdowns is significant at the five percent level at all values of *Hegemonic power volatility*. Similarly, with the exception of countries experiencing a growth rate above about 18 percent – which represents less than one percent of our sample – the effect of political shocks on authoritarian reversals is always significant at the five percent level.

We conduct two types of tests to show that adding the interaction term between economic and political shocks does not improve over the models presented in Table 4 and 5 of the main text. First, likelihood-ratio tests suggest that adding the interaction term does not improve the fit of the model. The test statistic (which follows a chi-squared distribution) comparing models 1 of Tables 5 and A33 is 0.148.²

Second, based on model 1 of Table A33, we calculate the marginal effect of *Growth rate* on authoritarian breakdowns when *Hegemonic power volatility* is set at its mean along with a 95 percent confidence interval. We also calculate the marginal effect of *Growth rate* when *Hegemonic power volatility* is one standard deviation above its mean. Again, we compute a 95 percent confidence interval. We find that the two confidence intervals overlap. In fact, 99 percent of the confidence interval computed when *Hegemonic power volatility* is at its mean is contained within the confidence interval calculated when *Hegemonic power volatility* is one standard deviation above its mean. Similarly, 90.8 percent of the latter confidence interval is contained within the former confidence interval. Therefore, the marginal effect of *Growth rate* on authoritarian breakdowns when *Hegemonic power volatility* is set at its mean is not statistically different than its marginal effect when *Hegemonic power volatility* is one standard deviation above its mean. This suggests that the effect of *Growth rate* is not

²To be statistically significant at the five percent level, the test statistic would need to be at least 5.991.

conditional on *Hegemonic power volatility*. This is also easy to see from the top left panel of Figure A1, since the confidence intervals are largely overlapped across different *Hegemonic power volatility* values.

We have conducted the same exercise to test whether the marginal effect of *Hegemonic power volatility* is conditional on *Growth rate*. Again, we find that the marginal effect of *Hegemonic power volatility* on authoritarian breakdowns when *Growth rate* is at its means is not statistically different than its marginal effect when *Growth rate* is one standard deviation above its mean. This time 85.2 percent of the confidence interval calculated when *Growth rate* is at its mean is contained within the confidence interval when *Growth rate* is one standard deviation above its mean. Likewise, 92.5 percent of the latter confidence interval is contained within the former confidence interval.

3 Spatial Probit Model

Some scholars recommend using a spatial econometric model to test diffusion of democratization (e.g., [Leeson and Dean 2009](#)). For example, in a spatial autoregressive model (i.e., $Y = \rho WY + X\beta + u$), a weighting matrix W is designed to capture a country's neighbors, and the model tests if a country's decision to democratize is influenced by its neighbors' level of democracy (i.e., if ρ is positive and statistically significant). Since we use a binary dependent variable, we employ a spatial autoregressive probit model to examine diffusion of democratization as a robustness test.³ That is, we test if ρ is positive and statistically significant in $Y^* = \rho WY^* + X\beta + u$, where Y^* is a latent variable.

We implement this spatial probit model using the `gmmprobit` command from the `McSpatial` package in R ([McMillen 2013](#)). Since we only include countries that begin the year as autocracies, our estimation captures whether democratization among neighbors within a given year induces democratization at home during that same year. Models of Table 4 in the main text are replicated,⁴ and the results are shown in column 1 of Table A34.⁵ As clearly demonstrated, the estimated ρ is $-.020$ and is statistically insignificant.⁶ Put differently, our spatial probit model suggests that contrary to the prediction of diffusion of democratization, autocracies are not more likely to democratize when their neighbors democratize during the same year. This finding echoes our findings in the main text. Column 2 redoes model 1

³For instance, see [LeSage and Pace \(2009\)](#) for a discussion of spatial autoregressive probit models.

⁴Since a spatial model does not need to include an independent variable to measure diffusion, it has less missing observations than the models in Table 4.

⁵When the oil income variable is employed, this spatial probit model fails to converge because this variable is highly skewed. Instead, we use a dichotomous oil variable taken from [Przeworski et al. 2000](#), which is equal to 1 if the average ratio of fuel exports to total exports in 1984-86 was greater than 50 percent, and 0 otherwise.

⁶This variable is labeled $WX\beta$ to denote WY^* .

without the control variable for the proportion of countries in the world that are democratic. Results are unchanged. In addition, the findings of other independent variables in Table A34 are consistent with our earlier results.

Table A34: Spatial Probit Estimation of Effect of Diffusion on Democratization

| | (1) | (2) |
|----------------------------|--------------------|--------------------|
| $WX\beta$ | -.020 (.069) | -.035 (.061) |
| % Dem. world | 1.610 (.406)*** | |
| Hegemonic power volatility | .149 (.067)** | .192 (.068)*** |
| Growth rate | -.026 (.011)** | -.031 (.012)*** |
| GDP pc (logged) | .260 (.066)*** | .314 (.068)*** |
| Oil | -.954 (.352)*** | -.963 (.327)*** |
| Muslim | -.001 (.002) | .001 (.002) |
| Catholic | .002 (.002) | .002 (.002) |
| Protestant | .003 (.003) | .003 (.003) |
| Ethnic frac. | .001 (.002) | .003 (.002) |
| Religious frac. | -.002 (.003) | -.002 (.002) |
| Brit. col. | .012 (.138) | .089 (.135) |
| # Past trans. | .262 (.058)*** | .286 (.059)*** |
| N | 6477 | 6477 |

Note: Spatial probit estimation ($Y^* = \rho WY^* + X\beta + u$). Standard errors in parentheses. All explanatory variables are lagged. *** $p < .01$, ** $p < .05$ and * $p < .1$.

We did not use spatial probit models as our main estimation technique because we believe the effect of diffusion occurs with at least one period lag. Put differently, neighbors' lagged level of democracy is more appropriate for the test of diffusion. Theoretically, it is difficult to justify a concurrent effect because it takes time for a country to observe and learn from its neighbors. Empirically, the yearly data does not offer information on the order of democratization when two or more countries democratize in the same year. Due to this data limitation, it is likely a later democratization is mistakenly assumed to have an effect on an earlier transition. Because of the temporal lags, the TSCS spatial autoregressive probit model can be estimated by a probit model (Beck, Gleditsch and Beardsley 2006).⁷ The

⁷In addition, we assume the error terms are temporally independent in the spatial au-

implication for our test of the diffusion of democratization is that instead of using a spatial autoregressive probit model, one can simply introduce a variable, like ours, to measure diffusion.

autoregressive probit model.

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