**Etable 1. Logistic regression models testing the effect of juvenile justice history on adult outcomes.**

|  |  |  |  |  |
| --- | --- | --- | --- | --- |
| **Outcomes** | **N(%)** | **OR** | **CI** | **P value** |
| Criminal records |   |   |   |   |
|  Misdemeanor charge | 463(31.0) | 1.7 | [0.8 , 3.3] | 0.1500 |
|  Felony charge | 133(7.7) | 2.8 | [1.2 , 6.7] | 0.0208 |
|  Violent crime felony charge | 87(5.8) | 2.6 | [1. , 8.7] | 0.0021 |
|  |  |  |  |  |
| Self-reported |   |   |   |   |
|  Recent police contact | 212(12.4) | 3.3 | [1.6 , 6.8] | 0.0014 |
|  Assault | 44(2.6) | 3.3 | [1.3 , 11.0] | 0.0365 |
|  Jail | 188(8.1) | 3.9 | [2.0 , 7.7] | 0.0007 |

Odds ratios combined from 10 multiply-imputed datasets with Rubin’s formula. Models adjusted survey weights and all covariates.

**Etable 2. Heterogeneity in ATT among people with juvenile justice history based on trimming and IPW adjusted models.**

|  |  |  |
| --- | --- | --- |
| Outcomes | Probation/Service/Treatment | Detention/Prison |
| OR | CI | p value | OR | CI | P value |
| Criminal records |   |   |   |   |   |   |
|  Misdemeanor charge | 1.3 | [0.7, 2.3] | 0.4442 | 4.6 | [1.8,11.8] | 0.0014 |
|  Felony charge | 2.1 | [1.0, 4.5] | 0.0547 | 6.4 | [2.3,18.4] | 0.0005 |
|  Violent crime felony charge | 2.4 | [1.0, 5.6] | 0.0510 | 7.2 | [2.6,19.9] | 0.0001 |
|  |  |  |  |  |  |  |
| Self-reported |   |   |   |   |   |   |
|  Recent police contact | 2.7 | [1.4,5.2] | 0.0042 | 1.8 | [0.8,9.8] | 0.2427 |
|  Assault | 2.2 | [0.8,6.2] | 0.1377 | 1.0 | [0.5,20.0] | 0.9685 |
|  Jail | 2.6 | [1.2,5.4] | 0.0112 | 6.3 | [2.0,19.8] | 0.0001 |

Odds ratios combined from 10 multiply-imputed datasets with Rubin’s formula. The study sample and ATT weights are the same as in Table 2.

**Etable 3. Logistic regression models testing the effect of juvenile justice history (non-residential v. no history and residential v. no history) on adult outcomes**

|  |  |  |
| --- | --- | --- |
|   | **non-residential v. no history** | **residential v. no history** |
| **OR** | **CI** | **P value** | **OR** | **CI** | **P value** |
| **Outcomes** |   |   |   |   |   |   |
| Criminal records |   |   |   |   |   |   |
|  Misdemeanor charge | 1.4 | [0.6 , 3.0] | 0.4008 | 2.9 | [1.1 , 7.6] | 0.0283 |
|  Felony charge | 2.2 | [0.8 , 5.6] | 0.1091 | 5.8 | [1.8 , 18.9] | 0.0032 |
|  Violent crime felony charge | 2.3 | [0.9 , 6.0] | 0.0920 | 4.4 | [1.2 , 16.4] | 0.0272 |
|   |   |   |   |   |   |   |
| Self-reported |  |   |   |   |   |   |
|  Recent police contact | 2.9 | [0.8 , 8.0] | 0.0123 | 1.4 | [0.5 , 3.9] | 0.5024 |
|  Assault | 2.8 | [1.1 , 11.6] | 0.0856 | 0.5 | [0.0 , 8.4] | 0.6393 |
|  Jail | 2.7 | [1.2 , 6.0] | 0.0126 | 4.7 | [1.7 , 12.7] | 0.0023 |

Odds ratios combined from 10 multiply-imputed datasets with Rubin’s formula. Models adjusted survey weights and all covariates.

**Efigure1. Ascertainment of the original Great Smoky Mountains study sample**



**Efigure 2.** **Covariate balance plot for probation/treatment/service vs. no history in childhood using no weights and IPW with different trimming thresholds for ATT.**

****

AI=American Indian; AA=African-American; SES=socioeconomic status; Dys=dysfunction; Sx=symptoms; Ins=insurance. “d” indicates trimming thresholds. SMD threshold at 0.10 is marked with a dashed line. d=0.0175 enables all covariates to have SMD below 0.10. Results are average of ten multiply-imputed datasets. The trimmed data have 166 (165-166) observations in the probation/treatment/service group and 810 (804-813) observations in the no history group.

**Efigure 3.** **Covariate balance plot for detention/jail vs. vs. no history in childhood using no weights and IPW for ATT.**



AI=American Indian; AA=African-American; SES=socioeconomic status; Dys=dysfunction; Sx=symptoms; Ins=insurance. Variable “Trauma” in log scale. “d” indicates trimming thresholds. SMD threshold at 0.10 is marked with a dashed line. Results are average of ten multiply-imputed datasets. The propensity score model for residential v. no history used no trimming and all the 57 observations in the residential involvement group are reserved.