Supplementary Information for:
Optimizing the Measurement of Sexism in Political Surveys

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1 Lucid Surveys

The main survey for this project was fielded July 31st - August 1st, 2019. Respondents were provided by Lucid, a firm that aggregates online panelists from a variety of different firms to sell a balanced sample of respondents to clients. A total of 1,207 unique respondents completed the questionnaire for the study. However, I dropped 91 respondents who progressed through the survey too quickly to realistically be reading the questions and an additional 14 respondents who gave inconsistent answers on a question about their age. After filtering out these respondents, I was ultimately left with a sample of 1,103 American adults.

The survey used to evaluate whether respondents viewed various items as being more or less political was fielded on May 26th, 2020. Respondents for this survey were also provided by Lucid and respondents who failed two out of three attention check items were filtered out of the sample. The ultimate sample size for this survey was 833 American adults.

Table SI.1 shows how the Lucid surveys compare to the 2018 Cooperative Congressional Election Study (using the weights provided by the CCES). Overall, both Lucid samples approximate the traits of the national adult population as characterized by the CCES. Perhaps most notably, the Lucid samples have very good partisan and ideological balance.

Table SI.1: Demographic composition of Lucid samples, compared to 2018 ACS

Trait	2018 CCES	2019 Lucid Survey	2020 Lucid Survey
Female	52%	53%	53%
Non-hispanic white	68%	73%	74%
College degree	30%	30%	43%
Democrats	44%	45%	41%
Republicans	38%	38%	43%
Liberals	28%	27%	25%
Conservatives	35%	30%	35%
Average age	47.7	46.1	48.5

2 Complete wording of Ambivalent Sexism Inventory Items

Respondents were asked to indicate their agreement or disagreement with each item on a six-point scale: disagree strongly, disagree somewhat, disagree slightly, agree slightly, agree somewhat, agree strongly.

Hostile sexism statements:

1. Many women are actually seeking special favors, such as hiring policies that favor them over men, under the guise of asking for equality.

2. Most women interpret innocent remarks or acts as being sexist.

3. Women are too easily offended.

4. Feminists are not seeking for women to have more power than men.

5. Most women fail to appreciate fully all that men do for them.

6. Women seek to gain power by getting control over men.

7. Women exaggerate problems they have at work.

8. Once a woman gets a man to commit to her, she usually tries to put him on a tight leash.

9. When women lose to men in a fair competition, they typically complain about being discriminated against.

10. There are actually very few women who get a kick out of teasing men by seeming sexually available and then refusing male advances.

11. Feminists are making entirely reasonable demands of men.

Benevolent sexism statements:

1. No matter how accomplished he is, a man is not truly complete as a person unless he has the love of a woman.

2. In a disaster, women ought not necessarily be rescued before men.

3. People are often truly happy in life without being romantically involved with a member of the other sex.

4. Many women have a quality of purity that few men possess.

5. Women should be cherished and protected by men.

6. Every man ought to have a woman whom he adores.

7. Men are complete without women.

8. A good woman should be set on a pedestal by her man.

9. Women, compared to men, tend to have a superior moral sensibility.

10. Men should be willing to sacrifice their own well being in order to provide financially for the women in their lives.

11. Women, as compared to men, tend to have a more refined sense of culture and good taste.

3   Complete wording of the Modern Sexism items

The wording for these items was taken directly from the VOTER survey. Here, respondents indicated agreement or disagreement on a four-point scale: strongly agree, somewhat agree, somewhat disagree, strongly disagree. The four statements:

1. Women often miss out on good jobs because of discrimination.

2. Women who complain about harassment often cause more problems than they solve.

3. Sexual harassment against women in the workplace is no longer a problem in the United States.

4. Increased opportunities for women have significantly improved the quality of life in the United States.

4   Structural Equation Model Accounting for Reverse-Coded Items

In Table 1 of the article, I show factor loadings from a maximum-likelihood factor analysis. The results from this model show that reverse-coded items perform particularly poorly, typically producing the smallest loading. In this section, I estimate a structural equation model that (1) incorporates a factor to account for whether an item is reverse-coded and (2) specifies which factor each manifest item should load on. The standardized coefficients from this model are presented in Table SI.2. Accounting for the reverse-coded items with the "methods factor" does help to increase the degree to which those reverse-coded items load on the expected factors. In some cases, the increase in these coefficients is considerable over what the standard factor analysis finds. But even after incorporating this methods factor into the analysis, the reverse-coded items still tend to load at a much lower level than most of the non-reverse-coded items.

## Table SI.2: Standardized Loadings from Structural Equation Model

	Hostile	Benevolent	Reversed
Hostile sexism items			
Many women are seeking special favors under guise of equality (hs1)	0.740		
Most women interpret innocent remarks or acts as being sexist (hs2)	0.759		
Women are too easily offended (hs3)	0.764		
Feminists are not seeking for women to have more power than men* (hs4)	-0.738		0.776
Most women fail to appreciate fully all that men do for them (hs5)	0.659		
Women seek to gain power by getting control over men (hs6)	0.785		
Women exaggerate problems they have at work (hs7)	0.757		
Once a woman gets a man to commit, she puts him on a tight leash (hs8)	0.683		
When women lose to men, they typically complain about discrimination (hs9)	0.733		
There are actually very few women who get a kick out of teasing men* (hs10)	-0.343		0.433
Feminists are making entirely reasonable demands of men* (hs11)	-0.607		0.722
Benevolent sexism items			
A man is not truly complete unless he has the love of a woman (bs1)		0.636	
In a disaster, women ought not necessarily be rescued before men* (bs2)		-0.376	0.569
People are often truly happy without a member of the other sex* (bs3)		-0.269	0.424
Many women have a quality of purity that few men possess (bs4)		0.670	
Women should be cherished and protected by men (bs5)		0.596	
Every man ought to have a woman whom he adores (bs6)		0.640	
Men are complete without women* (bs7)		-0.608	0.580
A good woman should be set on a pedestal by her man (bs8)		0.584	
Women, compared to men, tend to have a superior moral sensibility (bs9)		0.527	
Men should be willing to sacrifice for the women in their lives (bs10)		0.610	
Women tend to have a more refined sense of culture and good taste (bs11)		0.578	
Modern sexism items			
Women often miss out on good jobs because of discrimination* (m1)	-0.489		0.345
Women who complain about harassment cause more problems than solve (m2)	0.472		
Sexual harassment against women in the workplace is no longer a problem (m3)	0.314		
Increased opportunities for women have improved quality of life* (m4)	-0.315		0.125

Note: Entries are standardized coefficients from a structural equation model estimated using the sem command in Stata 15. N = 1,103. For complete wording of items see the appendix. *reverse coded item.

5 Testing for non-linearity in conditional effects

In the paper, I condition the treatment effects for the sex of the boss/candidate on the two sexism scales extracted from the factor analysis. However, these regression models impose a linear functional form to those treatment effects. Here, I use the approach recommended by Hainmueller, Mummolo and Xu (2018) to evaluate this linearity assumption. Figure SI.1 plots the treatment effects of gender across values of the hostile/modern sexism scale. The plot shows both the linear modeling of the marginal treatment effect as well as plotting the treatment effect for the bottom, middle, and top terciles of hostile sexism. The plot on the left shows the results for the boss conjoint – subjects who were below the mean on hostile sexism were more likely to choose bosses who had female names compared to those who had male names. Here, the tercile effects track very closely with the effects from a linear functional form and the Wald test produces a p-value of .40, indicating that we cannot be confident that the three-bin model is statistically distinct from the linear model.

Figure SI.1: Treatment effect of boss/politician gender conditioned by hostile sexism
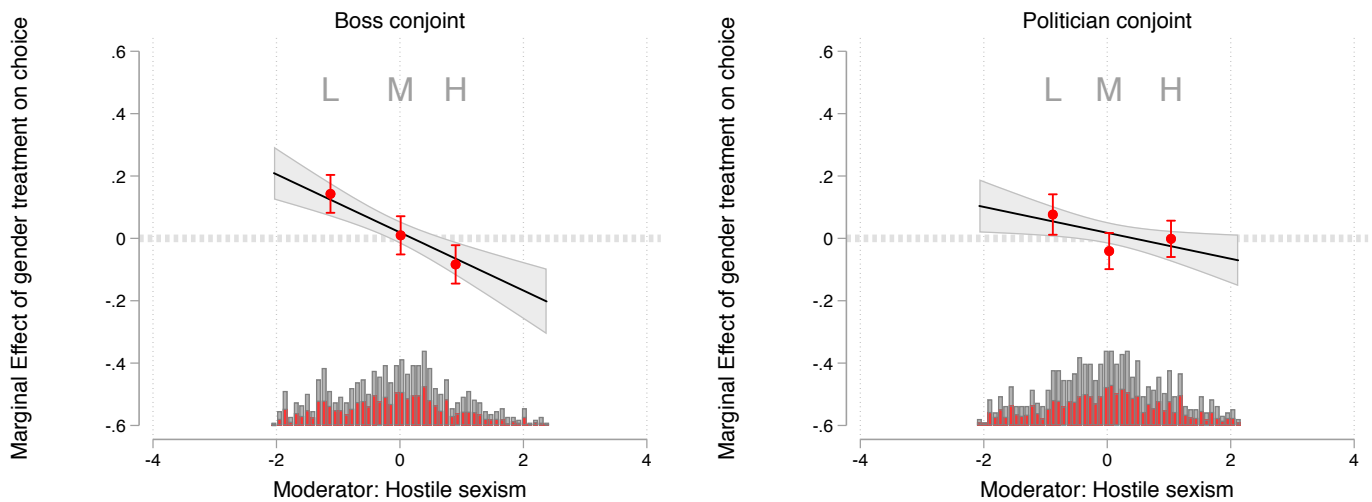
The plot on the right shows the influence of hostile sexism on moderating the gender treatment in the politician conjoint. In this plot, there appears to be a bit more departure from the linearly estimated effects. Nevertheless, the p-value for the Wald test is .27, meaning we cannot be confident that the three-bin model is distinct from the linear model.

6   Testing for non-linearity in conditional effects

In Table 3 of the paper, I show how hostile and benevolent sexism conditions the female treatment in the conjoint experiment. Here, I expand on the third model from that table and estimate a model that conditions the treatment effect not just on the sexism scales, but also on the respondent's age, gender, education, race, ideology, and partisanship. Even once accounting for all of these other factors, hostile sexism still conditions the female treatment effect with a very similar coefficient on the interaction term (-.063 in this model compared to -.069 in Table 3).

Table SI.3: Conjoint treatment effects controlling for other variables

	Coefficient
Female name treatment	0.138*
	(0.066)
Hostile Sexism	0.031***
	(0.008)
Female treatment × Hostile sexism	-0.064***
	(0.015)
Benevolent sexism	-0.003
	(0.007)
Female treatment × Benevolent sexism	0.009
	(0.015)
Age	0.001*
	(0.000)
Female treatment × Age	-0.002*
	(0.001)
Female respondent	0.007
	(0.013)
Female treatment × Female respondent	-0.013
	(0.026)
Ideology	0.006
	(0.006)
Female treatment × Ideology	-0.012
	(0.011)
Partisanship	0.008
	(0.008)
Female treatment × Partisanship	-0.017
	(0.015)
Education	-0.006
	(0.007)
Female treatment × Education	0.014
	(0.015)
White	-0.005
	(0.015)
Female treatment × White	0.012
	(0.031)
Intercept	0.429***
	(0.034)
N	6,618

Clustered standard errors in parentheses
* $p < 0.05$, ** $p < 0.01$, *** $p < 0.001$

7 Testing associations between sexism scales and opinions of politicians/policies

In the paper, I use the conjoint experiment as my main test of predictive validity. However, here I present additional tests based on correlational analysis between the sexism scales and political outcomes of interest. Specifically, I collected data on (1) how subjects evaluated prominent politicians and (2) the positions that subjects took on various issue items. For the first task, I included two women and two men who were running for the Democratic Party's presidential nomination at the time the survey was fielded (Kamala Harris, Elizabeth Warren, Bernie Sanders, and Joe Biden) as well as President Donald Trump and former Republican presidential candidate Mitt Romney. Respondents rated each candidate on a 0 to 100 scale with the end points labeled "very unfavorable" and "very favorable."

For the second task, respondents indicated their support for six policy proposals on a four-point scale ranging from "strongly oppose" to "strongly support." The policy proposals were:

1. Require corporations to show that they're not engaging in gender-based pay discrimination and fine companies that fail to close their gender-based pay gaps.

2. Allow schools, colleges, and universities to provide different athletic opportunities to girls and boys.

3. Require a minimum amount of renewable fuels (wind, solar, and hydroelectric) in the generation of electricity even if electricity prices increase somewhat.

4. Increase prison sentences for felons who have already committed two or more serious or violent crimes.

5. Spend $305 Billion in federal funds to repair and expand highways, bridges, and transit over the next 5 years.

6. Recognize Jerusalem as the capital of Israel and move the U.S. embassy from Tel Aviv to Jerusalem.

The six policy proposals were presented to respondents in a single grid and the order of appearance was randomized for each respondent. I selected two issues that directly deal with equal treatment of women compared to men – the proposal to hold companies responsible for gender-based pay gaps and the policy that would allow for boys and girls to be given different athletic opportunities (which would be a change from current law under Title IX).

For both sets of analyses, I estimate OLS regression models and include the latent scales for hostile sexism and benevolent sexism as the key independent variables. These scales are the same as those used in the conjoint analysis presented in the paper (the first and second factor from the factor analysis in Table 1). The models also include controls for partisanship, ideology, age, education, race, and gender. The full model estimates can be found at the end of this Supplementary Information document.

7.1 Results

Here, I plot the coefficients (and 95% confidence intervals) for the hostile/modern (first factor) and benevolent sexism (second factor) scales extracted from the factor analysis presented in the paper. Figure SI.2 presents these results from the analysis of favorability ratings. Benevolent sexism is a statistically significant predictor in most of the models. However, it is noteworthy that the coefficient for benevolent sexism is positive for each politician, regardless of the party or gender of that politician.

Figure SI.2: Marginal effects of hostile and benevolent sexism on politician favorability ratings
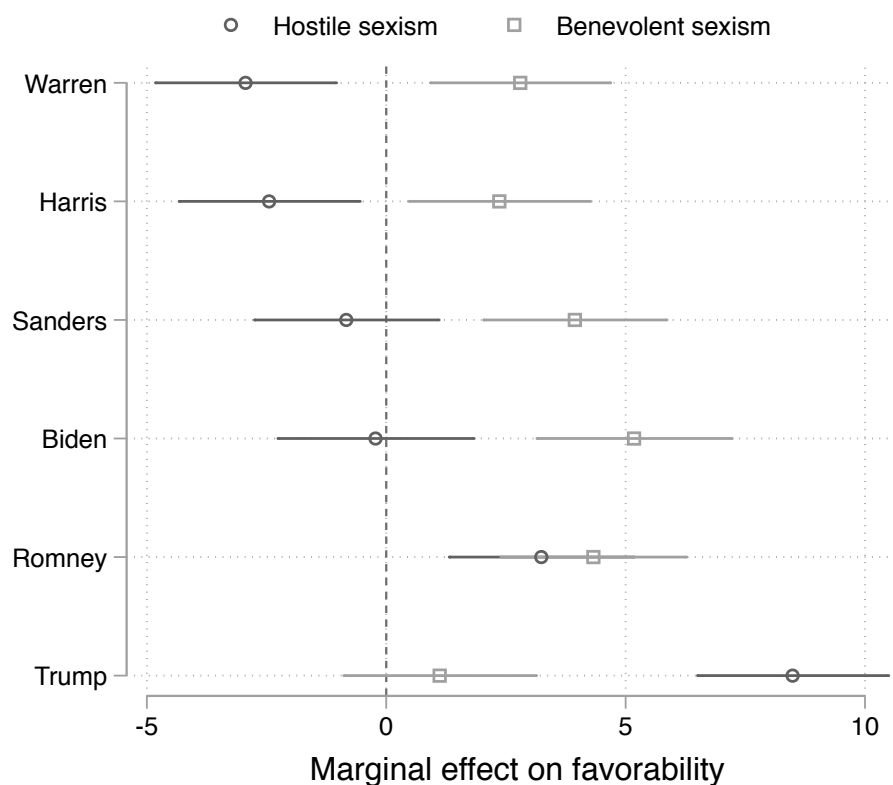
In contrast to the patterns for benevolent sexism, the association between hostile sexism and favorability ratings does appear to differ according to both the gender and the party of the candidate being evaluated. A one standard deviation increase in hostile sexism is associated with about a 2.5-3 point decrease in ratings of both Warren and Harris. For Sanders and Biden, the coefficient is close to zero and is not statistically significant. The coefficient for hostile sexism is positively signed and statistically significant in the models for ratings of the two Republican politicians. Hostile sexism is most strongly associated with

attitudes towards Trump – a one standard deviation increase in hostile sexism is associated with more than an 8 point increase in how favorably respondents rated him.

Figure SI.3 plots the coefficients for hostile and benevolent sexism from the models gauging support for six policies. For half of the policies, benevolent sexism has a coefficient that is small and not statistically significant. One notable exception is in the model for allowing different athletic opportunities for boys and girls. Here, a one standard deviation increase in benevolent sexism is associated with a nearly 5 point increase in support for the policy (which runs counter to the mandate in Title IX). Benevolent sexism is also positively associated with support for recognizing Jerusalem and for a policy that would fine companies that engage in gender-based pay discrimination.

Figure SI.3: Marginal effects of hostile and benevolent sexism on support for policies
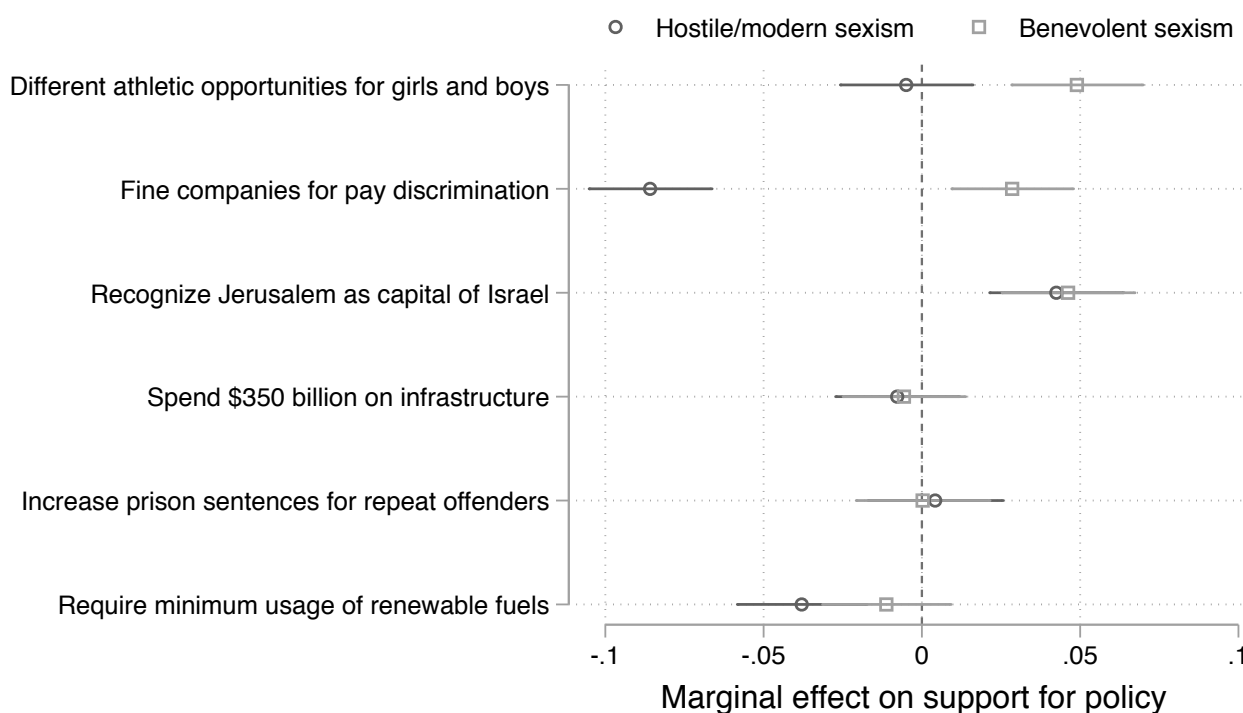
The coefficient for hostile sexism is small and not statistically significant in the model for offering different athletic opportunities to boys and girls, as well as in two other models. Hostile sexism is most strongly related to opposition to the policy that involves fining companies that have gender-related pay gaps. A one standard deviation increase in hostile sexism is associated with more than an 8 percentage point decline in support for such a

policy. Hostile sexism is also more modestly related to support for recognizing Jerusalem as the capital of Israel and opposition to requiring a minimum usage of renewable fuels.

8 Test-Retest Reliability

As a measure that is meant to capture an attitude that is likely to change slowly, if at all, it is useful to examine whether respondents answer hostile sexism items in a similar way over shorter and longer periods of time. While the Lucid survey was cross-sectional, hostile sexism items have been fielded on two different panel surveys. Both panel surveys were conducted by YouGov:

1. The first wave of the survey was conducted October 25th-31st, 2016; the second wave was fielded March 1st-20th, 2017, and the third wave was in the field from July 17th-August 3rd, 2018. In this paper, I limit the analysis to the 731 respondents who answered the survey in each wave. Respondents were recruited and interviewed by the survey firm YouGov. Respondents were asked four hostile sexism items at three different points across three years – October 2016, March 2017, and July 2018. The value of this survey is that respondents were also asked items from the child rearing scale used to measure authoritarianism.

2. The first wave of the survey was a nationally representative sample of 2,118 American adults fielded from September 25 to September 26, 2018. The recontact wave of the survey began fielding on October 7th. After cleaning, there were 1,728 respondents who took both surveys. Respondents were recruited and interviewed by the survey firm YouGov. This sample answered 6 hostile sexism items about one month apart – September 2018 and October 2018.

Table SI.4 presents the intraclass correlation coefficients for each item that was asked in each of the panel surveys. Intraclass correlation coefficients indicate how correlated each item is within each respondent over time. The bottom of the table presents intraclass correlation coefficients for four items from the child rearing authoritarian scale as a point of comparison. Correlation coefficients only appear for the dataset in which the item was asked. For example, six hostile sexism items were asked in the 2018 YouGov panel survey, two of those items were also on the longer YouGov panel survey, and the longer YouGov panel also included two items not on the short panel. The authoritarian items were only included on the long panel.

Overall, Table SI.4 reveals reasonably high intraclass correlations for most of the hostile sexism items. In fact, nearly every hostile sexism item has retest reliability that is higher than that for the child rearing items. The one exception to this pattern is the reverse coded item ("There are actually very few women who get a kick out of teasing men"). The correlation for this item is below .5, possibly owing to issues that respondents tend to have with reverse coded items in general. However, it is worth nothing that the other reverse coded item has a fairly strong intraclass correlation (r = .668).

8.1 Intraclass Correlation Coefficients for Modern Sexism Items

To determine the test/retest reliability of the modern sexism items that I analyze in the paper, I use the VOTER survey panel. The VOTER survey is administered by YouGov and

Table SI.4: Intraclass Correlation Coefficients

Item	3-wave panel (2016-2018)	2-wave panel (2018)
Hostile sexism items		
Women seek to gain power by getting control over men	.659	.689
Women are too easily offended	.705	.743
Many women are actually seeking special favors...	.703	
When women lose to men in a fair competition, they typically...	.643	
Most women fail to appreciate fully all that men do for them		.698
Most women interpret innocent remarks or acts as being sexist		.663
Feminists are not seeking for women to have more power than men*		.668
There are actually very few women who get a kick out of teasing men*		.429
Authoritarian items		
Independence vs. respect for elders	.582	
Good manners vs. curiosity	.541	
Self reliance vs. obedience	.627	
Being considerate vs. well behaved	.472	
N	720	1,439

Note: *reverse coded items.

is designed to be nationally representative of American adults. Respondents were asked the sexism items in December, 2016 and again in May, 2018. Table SI.5 shows the intraclass correlation coefficients for each of the four modern sexism items as well as for an additional item that is nearly identical to one item from the hostile sexism scale. While the ICC for the first item is reasonably high, they are quite low for the last two items. Notably, the one item included that is derived from the hostile sexism scales produces the highest ICC.

Table SI.5: Intraclass Correlation Coefficients

Item	ICC
Modern sexism items	
Women often miss out on good jobs because of discrimination.	.680
Women who complain about harassment often cause more problems than they solve.	.603
Sexual harassment against women in the workplace is no longer a problem in the U.S.	.570
Increased opportunities for women have significantly improved the quality of life in the U.S.	.376
Hostile sexism item	
When women demand equality these days, they are actually seeking special favors.	.724
Old fashioned sexism item	
Women should return to their traditional roles in society.	.682
N	4,626

9 Hostile sexism and sex-based prejudice by party of politician

Respondents who were randomized into the politician conjoint experiment were also randomly assigned to view profiles that were either all Democratic politicians or all Republican politicians. Table SI.6 reproduces the model from Table 3, but dividing the models based on whether respondents saw Democratic or Republican profiles. Notably, hostile sexism strongly conditioned the effects of the female treatment when the profiles were Democrats, but not when they were Republicans.

Table SI.6: Models estimating treatment effect of politician sex conditional on hostile and benevolent sexism, by party of politician

	Democrats	Republicans
Female treatment	0.005	0.030
	(0.024)	(0.026)
Hostile sexism	0.036*	0.007
	(0.014)	(0.014)
Female treatment x hostile sexism	-0.070*	-0.014
	(0.028)	(0.029)
Benevolent sexism	-0.014	0.011
	(0.014)	(0.014)
Female treatment x benevolent sexism	0.030	-0.025
	(0.027)	(0.031)
Intercept	0.497***	0.486***
	(0.012)	(0.013)
N	1704	1692

Standard errors in parentheses

* $p < 0.05$, ** $p < 0.01$, *** $p < 0.001$

10 Hostile sexism items oriented towards men

In a critique of current scholarship focusing on sexist attitudes, Zigerell (2019) argues that it is a problem that "gender-attitudes items did not measure attitudes disfavoring men" (p. 720). Of course, it is easy to see why items that ask about racism and sexism tend to focus on views toward the group that has traditionally been the subject of discrimination and prejudice. Nevertheless, in the Lucid survey, I also included five hostile sexism items that I re-wrote in order to be directed towards men:

1. When men lose to women in a fair competition, they typically complain about being discriminated against.

2. Many men are actually seeking special favors, such as hiring policies that favor them over women, under the guise of asking for "equality."

3. Men seek to gain power by getting control over women.

4. Most men interpret innocent remarks or acts as being sexist.

5. Men are too easily offended.

When I scale these items using an IRT graded response model, the resulting latent scale is only modestly correlated with hostile sexism (r = .19) and benevolent sexism (r = .25). Thus, it is highly unlikely that failure to account for resentment and prejudice towards men is biasing the findings regarding the role of sexist attitudes on opinions and vote choice. Indeed, including this scale in the models of candidate favorability or issue attitudes does not alter the patterns of results for the hostile sexism scale. The male-oriented scale demonstrates no association with gender-related policies, with coefficients close to zero and p-values above .95 in the models asking about support for closing the gender pay gap and relaxing Title IX. Additionally, as shown in Figure SI.4, the male-oriented hostile sexism scale provides no real predictive power on whether a subject chose a female versus a male boss or politician in the conjoint experiment. Thus, it is not a predictor of sex-based prejudice.

Figure SI.4: Treatment effect of boss/politician gender conditioned by male-oriented hostile sexism items
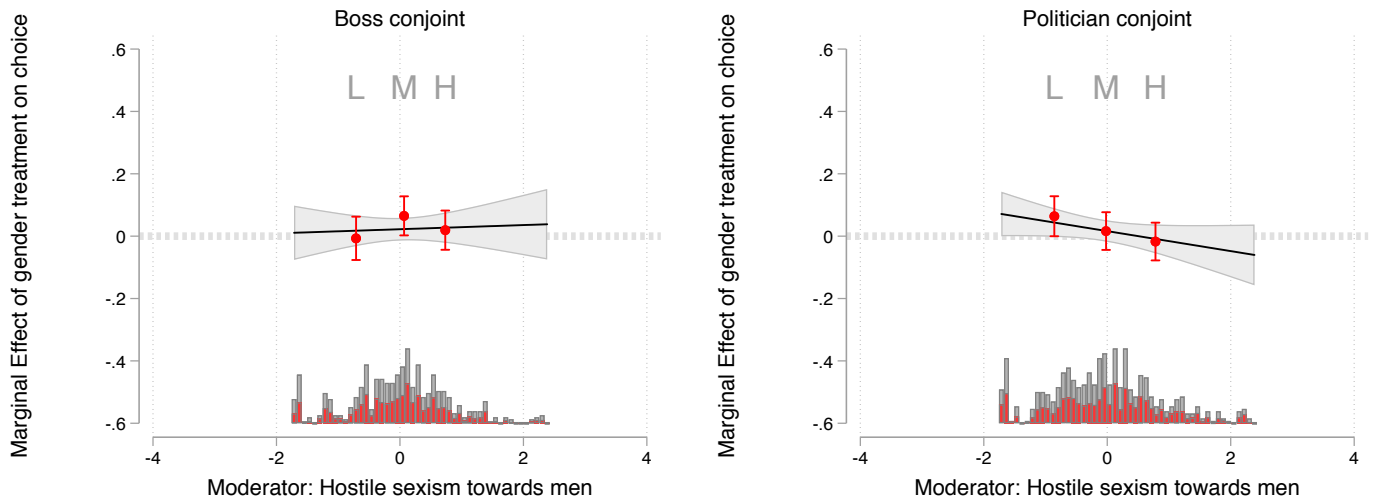
Overall, then, the traditional hostile sexism scale accurately captures resentment and prejudice towards women and capturing similar sentiments towards men appears to be unnecessary. Indeed, hostility towards men does not appear to predict sex-based prejudice or

11 Full models for analysis of favorability ratings and policy support

In this section, I provide the full information for the regression models used to produce Figures SI.2 and SI.3.

Table SI.7: Full results from models on politician favorability

	Trump	Sanders	Biden	Warren	Harris	Romney
Hostile sexism	8.487***	-0.835	-0.224	-2.446*	-2.939**	3.239***
	(1.014)	(0.978)	(1.040)	(0.960)	(0.960)	(0.982)
Benevolent sexism	1.117	3.939***	5.176***	2.362*	2.799**	4.326***
	(1.021)	(0.971)	(1.035)	(0.962)	(0.953)	(0.985)
Party: Ind/other	14.164***	-20.598***	-22.378***	-17.614***	-15.454***	-3.571
	(2.741)	(2.610)	(2.775)	(2.570)	(2.562)	(2.631)
Party: Republican	44.568***	-27.034***	-26.165***	-20.266***	-19.858***	5.751*
	(2.399)	(2.303)	(2.457)	(2.277)	(2.272)	(2.315)
Ideology: Liberal	-2.116	-7.754*	-3.601	-7.905*	-7.845*	0.751
	(3.461)	(3.253)	(3.469)	(3.187)	(3.191)	(3.344)
Ideology: Moderate	2.774	-15.605***	-6.656*	-21.914***	-18.862***	4.312
	(3.180)	(2.992)	(3.190)	(2.935)	(2.940)	(3.086)
Ideology: Conservative	10.819**	-25.507***	-14.080***	-28.161***	-25.290***	8.032*
	(3.709)	(3.532)	(3.756)	(3.474)	(3.486)	(3.596)
Ideology: Very cons,.	20.300***	-29.501***	-21.251***	-35.013***	-30.759***	4.280
	(4.142)	(4.009)	(4.249)	(3.941)	(3.927)	(4.046)
Ideology: Not sure	-3.956	-19.757***	-16.769***	-23.573***	-21.389***	1.946
	(4.316)	(4.045)	(4.345)	(4.016)	(4.012)	(4.180)
Age: 31-44	7.253**	-0.032	5.398*	3.720	3.621	1.217
	(2.545)	(2.427)	(2.594)	(2.410)	(2.397)	(2.450)
Age: 45-64	5.903*	-8.814***	4.223	2.744	3.020	1.995
	(2.402)	(2.287)	(2.450)	(2.265)	(2.256)	(2.317)
Age: 65+	6.437*	-13.573***	3.487	-1.042	-2.477	-0.653
	(2.873)	(2.770)	(2.947)	(2.717)	(2.710)	(2.802)
Female	-0.236	-0.469	0.883	-0.293	-1.649	2.895
	(1.797)	(1.734)	(1.848)	(1.710)	(1.702)	(1.745)
Education: Some college	-2.846	-0.970	-0.627	0.374	1.573	-0.421
	(2.160)	(2.076)	(2.213)	(2.048)	(2.038)	(2.089)
Education: College	-5.741*	1.688	2.076	6.301*	4.735	2.676
	(2.615)	(2.526)	(2.686)	(2.485)	(2.465)	(2.536)
Education: Post-grad	-2.758	6.686*	8.519*	12.169***	15.229***	11.242***
	(3.272)	(3.188)	(3.382)	(3.154)	(3.119)	(3.202)
White	5.524**	-3.173	-1.029	-2.948	-5.353**	0.627
	(2.124)	(2.012)	(2.143)	(1.989)	(1.979)	(2.049)
Intercept	10.351**	87.229***	67.120***	73.676***	70.999***	28.652***
	(3.847)	(3.588)	(3.839)	(3.560)	(3.553)	(3.699)
N	1029	1078	1076	1059	1061	1062

Standard errors in parentheses

* $p < 0.05$, ** $p < 0.01$, *** $p < 0.001$

Table SI.8: Full results from models on policy support

	paygap	title9	renewable	sentencing	infrastructure	jerusalem
Hostile sexism	-0.086***	-0.005	-0.038***	0.004	-0.008	0.042***
	(0.010)	(0.011)	(0.010)	(0.011)	(0.010)	(0.011)
Benevolent sexism	0.029**	0.049***	-0.011	0.000	-0.006	0.046***
	(0.010)	(0.011)	(0.010)	(0.011)	(0.010)	(0.011)
Party: Ind/other	-0.062*	-0.024	-0.103***	0.001	-0.050	-0.028
	(0.026)	(0.028)	(0.028)	(0.029)	(0.026)	(0.029)
Party: Republican	-0.060**	-0.001	-0.112***	0.053*	-0.032	0.108***
	(0.023)	(0.025)	(0.024)	(0.025)	(0.023)	(0.025)
Ideology: Liberal	-0.037	0.028	0.008	0.055	0.039	-0.068
	(0.033)	(0.035)	(0.035)	(0.036)	(0.033)	(0.036)
Ideology: Moderate	-0.088**	0.024	-0.079*	0.103**	0.012	-0.005
	(0.030)	(0.033)	(0.032)	(0.033)	(0.030)	(0.033)
Ideology: Conservative	-0.125***	0.026	-0.088*	0.089*	-0.003	0.070
	(0.035)	(0.038)	(0.037)	(0.039)	(0.036)	(0.039)
Ideology: Very cons.	-0.112**	0.004	-0.096*	0.125**	-0.025	0.089*
	(0.040)	(0.043)	(0.042)	(0.044)	(0.040)	(0.044)
Ideology: Not sure	-0.022	0.038	-0.070	0.098*	-0.029	0.000
	(0.040)	(0.043)	(0.042)	(0.044)	(0.040)	(0.044)
Age: 31-44	0.008	0.053*	0.016	0.146***	0.053*	0.025
	(0.024)	(0.026)	(0.026)	(0.027)	(0.024)	(0.026)
Age: 45-64	0.027	0.035	-0.040	0.206***	0.119***	0.046
	(0.023)	(0.025)	(0.024)	(0.025)	(0.023)	(0.025)
Age: 65+	0.033	0.044	-0.041	0.177***	0.176***	0.072*
	(0.028)	(0.030)	(0.029)	(0.030)	(0.028)	(0.030)
Female	0.034*	0.013	-0.022	0.033	-0.072***	0.013
	(0.017)	(0.019)	(0.018)	(0.019)	(0.017)	(0.019)
Education: Some college	0.018	-0.030	-0.020	-0.010	0.000	-0.012
	(0.021)	(0.022)	(0.022)	(0.023)	(0.021)	(0.023)
Education: College	-0.000	-0.049	-0.029	-0.052	-0.014	-0.016
	(0.025)	(0.027)	(0.027)	(0.028)	(0.025)	(0.028)
Education: Post-grad	0.064*	-0.026	-0.004	-0.065	0.014	-0.054
	(0.032)	(0.035)	(0.034)	(0.035)	(0.032)	(0.035)
White	-0.012	-0.031	0.027	0.030	0.037	-0.007
	(0.020)	(0.022)	(0.021)	(0.022)	(0.020)	(0.022)
Intercept	0.794***	0.681***	0.795***	0.460***	0.641***	0.494***
	(0.036)	(0.039)	(0.038)	(0.040)	(0.036)	(0.039)
N	1103	1103	1103	1103	1103	1103

Standard errors in parentheses

* $p < 0.05$, ** $p < 0.01$, *** $p < 0.001$

References

Hainmueller, Jens, Jonathan Mummolo and Yiqing Xu. 2018. "How Much Should We Trust Estimates from Multiplicative Interaction Models? Simple Tools to Improve Empirical Practice." *Political Analysis* pp. 1–30.

Zigerell, L.J. 2019. "Left Unchecked: Political Hegemony in Political Science and the Flaws It Can Cause." *PS Political Science & Politics* 52(4):720–723.